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**A Founding Penalty:
Evidence from an audit study on gender, entrepreneurship, and future employment**

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There is both widespread interest in encouraging entrepreneurship and universal recognition that the vast majority of these founders will fail, which raises an important unanswered question: What happens to ex-founders when they apply for jobs? While existing research has identified many factors that facilitate movement out of an established organization and into entrepreneurship, far less attention has been devoted to understanding what transpires during the return journey. Most notably, how employers evaluate entrepreneurial experience at the point of hire. We propose that employers penalize job candidates with a history of founding a new venture because they believe them to be worse fits and less committed employees than comparable candidates without founding experience. We further predict that the discount for having been an entrepreneur will diminish when other stereotypes about the candidate, and in particular based on their gender, will contradict the negative beliefs about ex-founders. We test our proposition using a resume-based audit and an experimental survey. The audit reveals that founding significantly reduces the likelihood an employer interviews a male candidate, but there is no comparable penalty for female ex-founders. The experimental survey confirms the gendered nature of the founding penalty and provides evidence that it results from employers concerns that founders are less committed and worse organizational fits than non-founders. Critically, the survey also indicates that these concerns are mitigated for women, helping to explain why they suffer no equivalent founding penalty.

Introduction

In modern labor markets, career trajectories are changing as workers increasingly opt to transcend conventional boundaries of wage work and move to less typical employment relationships (e.g., Pedulla, 2020; Leung, 2014). Entrepreneurship is one example of such non-standard arrangements, with a significant share of individuals choosing to leave the stability of being a salaried worker in order to start their own business (Burton et al. 2002; Carnahan, Agarwal and Campbell, 2012; Kacperczyk, 2013; Vissa, 2011; 2012). Migration into entrepreneurship has been spurred along by financial incentives from state and federal governments (Armanios et al., 2017; Lanahan and Feldman, 2015) and new training ecosystems that have spread among universities and private incubators (Hallen, Cohen, and Bingham, 2020). Underlying these specific initiatives is a general notion that entrepreneurial activities are economically and socially desirable, acting as a catalyst of economic growth and prosperity (Haltiwanger et al., 2012; Malecki, 2018; Urbano, Aparicio, and Audretsch, 2018), and therefore we all benefit by encouraging more individuals to launch new firms (Feldman and Florida, 1994; Marx, 2011).

But while entrepreneurship may result in substantial gains for nations and regions, we still have only a partial understanding of whether similar rewards accrue to the individuals who decide to found a company. In examining the benefits and costs of entrepreneurial careers, past studies have paid disproportionate attention to *movements into* entrepreneurship, documenting the income lost at the point of entry or during the spell of entrepreneurship (e.g., Borjas and Bronars, 1989; Hall and Woodward, 2010; Hamilton, 2000; Moskowitz and Vissing-Jørgensen, 2002). But beyond founding and managing a new venture, the pursuit of entrepreneurship often involves a final stage: exit. And given prohibitively high failure rates (Dunne et al., 1988; Gompers et al., 2010; Mata and Portugal, 1994; Headd, 2003), a founder's future often involves returning to wage work (Baptista et al., 2012; Campbell, 2013, Luzzi and Sasson, 2016; Manso, 2016). Relatively little research, however, has considered founders' future prospects in wage work and even the early evidence from studies of future wages remain mixed (Campbell, 2013; Luzzi and Sasson, 2016; Manso, 2016) — limiting our accounting of the full costs of an entrepreneurial career. Indeed, Burton et al., (2017: 239) state that “Movements between entrepreneurship

and paid employment—including both transitions to entrepreneurship and re-entry into paid employment—are remarkably common, yet have only recently begun to attract scholarly attention.”

In this study, we extend this small and inconclusive line of inquiry by turning our attention to employers and their evaluations of ex-founders during recruitment. Given that many founders will, at some point, solicit established firms for employment as salaried workers, employers will inevitably function as a final but so-far unconsidered audience for ex-founders, with their evaluations having a pervasive effect on the founder’s career prospects upon re-entry. Because this assessment is part of a regular screening process of job candidates by employers, we leverage the demand-side perspectives on hiring and employers’ evaluations of atypical employment histories (e.g., Bills, Stasio and Gerxhani, 2017; Pedulla, 2016; 2020; Leung, 2014; Rivera, 2020) to develop a set of predictions about assessments of ex-founders, and the consequences of such evaluations for these individuals’ prospective return to paid employment. These theories posit that, in addition to inferring ability, employers seek evidence of a job candidate’s fit with the organization (Chatman 1989; O’Reilly et al. 1991; Rivera 2012) and their likelihood of commitment (e.g., Galperin et al 2019). Given that leaving to found a new venture can resemble a pattern of interrupted employment associated with fitness (Leung 2014; Pedulla 2016; 2020) and commitment-based penalties (Weisshaar 2018), we expect entrepreneurship to raise questions around a candidate’s willingness to commit to the organization or to fit in wage work, more generally. In short, we predict that becoming a founder elicits a range of negative stereotypes, including that the candidate could be difficult to manage, or retain, that they are not a team player, and that these perceptual penalties function to reduce an employer’s willingness to advance the candidate.

However, our theory does not describe a uniform penalty that applies to all ex-founders. Instead, to probe the proposed underlying mechanism, we further predict that the discount for having been an entrepreneur will diminish when other stereotypes about the candidate contradict the negative beliefs about ex-founders. In particular, we draw on research on stereotype contrasts (e.g., Blair 2002; Blair, Ma, and Lenton 2001; Dasgupta and Greenwald 2001) to propose that female and male ex-founders will have significantly different experiences at re-entry, and identify two reasons why gender might diminish the

penalty for female ex-founders. First, positive stereotypes about women as more compliant and more committed (e.g., Fiske et al., 2002; Heilman, 1980; 1984; Galperin et al., 2019; Rivera, 2017) can offset the specific negative stereotypes employers attach to ex-founders. Second, deeply held cultural beliefs that entrepreneurship is a male-typed activity (Heilman and Chen, 2003; Thébaud 2010, 2015b) will decrease the informational value of founding when the candidate is a woman, reducing the likelihood that employers associate female ex-founders with the initial commitment and fitness concerns. Critically, the centrality of perception in our account—the notion that a founding penalty derives from the interpretation of entrepreneurship rather than the act itself—implies a potential respite to groups traditionally disadvantaged in the entrepreneurial process. In sum, we expect that the employment-related costs of an entrepreneurial career are highly contingent on gender, leading to a potential reduction in the risk of re-entry challenges for female founders.

Methodologically, the endogeneity of an entrepreneurial transition poses a formidable challenge for causal estimates. To address these concerns, we adopt a resume-based audit study (Kang et al. 2016; Pager and Quillian 2005; Pager and Shepherd 2008; Rivera and Tilcsik 2016; Tilcsik 2011), in which we send artificial resumes to real job postings. With this research design, it is possible to randomize entry into entrepreneurship for otherwise identical candidates and thus assess the influence of pursuing entrepreneurship on employer interest net of any applicant-level differences. Between July 2017 and May 2018, we submitted 1,223 applications for two occupations in twelve cities across the United States and found evidence that recruitment chances are significantly lower for ex-entrepreneurs than they are for wage workers, but also that the likelihood of an interview is significantly higher for female than male ex-entrepreneurs. We complement this first study with an experimental survey of marketing hiring managers to examine how founding and gender jointly affect an employer's perception of an applicant's fitness. The results of this second experiment provide additional evidence that the founding penalty is motivated by employer concerns around fitness and commitment, but that these concerns are less strongly attached to female ex-founders.

THEORY

Returns to Entrepreneurship

Ample research to date has considered entrepreneurship, or the act of leaving paid employment to start a new venture and exploit market opportunities, as a career choice (Burton et al., 2017; Sørensen and Sharkey, 2014; Kacperczyk and Marx, 2016). Accordingly, scholars have increasingly investigated the costs that individuals bear when pursuing an entrepreneurial career (e.g., Hall and Woodward, 2010; Hamilton, 2001; Moskowitz and Vissing-Jørgensen, 2002). There has been growing recognition, however, that the extant research has paid little attention to the final stage of the entrepreneurial process: a founder's exit (Burton et al., 2017). For many, however, the return to paid employment will be the final stage of the entrepreneurial lifecycle, because the vast majority of founders will, at some point, abandon their entrepreneurial career and return to the workforce as paid employees (Burton et al., 2017; Campbell, 2012; Luzzi and Sasson, 2016; Baptista et al., 2012; Kacperczyk and Younkin, 2019; Sørensen and Phillips 2011), either because their venture failed or because the founders elect to re-enter paid work.

Although scant, the few extant studies of "returning founders" focus primarily on estimating the effect of founding on future wages (Campbell, 2013; Luzzi and Sasson, 2016; Manso, 2016), and fail to deliver consistent findings. On the one hand, a limited number of studies document positive returns for founders who sort back into paid employment (Luzzi and Sasson, 2016). For example, using a sample of early-stage start-up employees and matched non-start-up employees in the semiconductor industry, Campbell (2013) finds large and persistent earnings gains for founders compared to non-founders, with payoffs robust to whether the firm failed or succeeded. This positive view of entrepreneurial payoffs has been challenged by studies that fail to deliver similar findings. In a study of Portuguese tax records, for example, Baptista et al. (2012) find that former founders receive more senior positions than wage employees, but they also earn less than they would have earned, if they had remained wage employees. Other scholars have found similar results, with some variations among the conditions under which wage penalties are more or less prevalent (Bruce et al. 2004, Kaiser and Malchow-Møller 2011, Mahieu et al. 2018). More generally, the focus of prior inquiries has been on wage variation, but before any penalty or premium in pay can attach, employers must first evaluate whether an ex-founder is a desirable hire.

Employers thus function as a critical audience, and their appraisals of ex-founders determine the ex-founder's subsequent chances to re-enter the labor market. Because this influences the net sum of benefits ascribed to entrepreneurship, in what follows, we turn our attention to the employer's evaluation of ex-founders at the critical stage of re-entry into paid employment.

Evaluations of Ex-Entrepreneurs

Given that the demand-side employer's assessments of ex-founders can have a profound impact on the consequences of an entrepreneurial career, a fundamental question is how prospective employers will perceive "returning entrepreneurs" at the point of hire.¹ Because assessments of ex-founders are part of a broader hiring process, whereby employers screen out job candidates, we build on research pertaining to the demand-side approaches to hiring and theories of employer evaluations of non-standard employment (e.g., Bills et al., 2017; Leung, 2014; Pedulla, 2020) to develop a set of predictions about how employers perceive ex-founders and how their evaluations will, in turn, affect these job candidates' chances of hire.

Studies of screening in labor markets have long established that, in evaluating prospective hires, employers are attuned to two sets of signals: (1) evidence of the candidate's ability to perform at a high level (Bills, 1990; Rivera, 2012; Rivera and Tilcsik, 2016; Ferguson and Hasan, 2013), and (2) evidence of a cultural fit and willingness to commit to the future employer (Leung, 2016; Pedulla, 2016; 2020). While the relevance of the former can easily be anticipated, the case for the latter is equally understandable. In particular, employers recognize that organizations are not uniform but possess distinct cultural blueprints (Baron and Hannan, 2002) that employees must fit, not only in terms of their skills but also culturally (Rivera, 2012), if the firm and employee are to succeed (Burton and Beckman, 2007).² In

¹ Entrepreneurship can refer both to self-employment and the creation of a new organization (Levine and Rubinstein 2017). We focus our attention on the creation of a new venture rather than self-employment because the two are driven by different motives and may therefore trigger different perceptions among employers. Our theory therefore pertains only to instances where individuals created new organizations.

² An organizational culture may reflect formal structures (Baron and Hannan, 2002)—how people are promoted or assigned tasks—or shared norms and beliefs—for example, political preferences (Bermiss and McDonald 2018)—but irrespective of the specific form of culture, the findings are consistent: mismatches fail and exit.

addition, employers seek to identify which applicants are liable to commit to the firm, given the high search and training costs that come with voluntary turnover (e.g., Osterman, 1987; Delery and Shaw, 2001; Leung, 2012; Galperin et al., 2019). But fit and commitment are rarely observed directly, requiring that employers infer an applicant's attributes or preferences from the scant available signals found in their career history (Spence 1973; Leung 2012). For founders, entrepreneurship becomes embedded in this history (Burton et al. 2002; Sørensen and Sharkey, 2014; Kacperczyk and Marx, 2016), and therefore has the potential to function as one such signal, influencing employer assessments in the same fashion as other career transitions (Bills, 1990).

The question is whether founding a firm will influence an employer's assessment of either quality or fitness/commitment? There is a strong reason to anticipate that the manner in which founding will affect the perceived quality is liable to vary according to unobservable differences in the preferences of an employer and the nature of the position itself. This heterogeneity makes it unlikely that founding has a consistently positive or negative effect on the average employer's perception of their skillset. However, widely shared cultural beliefs that founders are particularly independent or averse to corporate authority will suggest a more consistently negative relationship between founding and perceptions of commitment and fitness. First and with respect to fit, founders are often believed to dislike bureaucratic authority (Wasserman, 2006; Hamilton, 2000; Benz, 2009; Hurst and Pugsley, 2011), and to be motivated by the need for independence and autonomy (Hamilton, 2000; Moskowitz and Vissing-Jørgensen, 2002). These stereotypes are readily available to inform prospective employers' perceptions, possibly leading to expectations that the candidate is unfit for wage work because they may be unwilling to acquiesce to a prospective boss, difficult/costly to manage or reluctant to be a member of a team.

Second and with respect to commitment, founders are commonly believed to possess entrepreneurial traits that conflict with working for someone else for any length of time (e.g., Kihlstrom and Laffont 1978; McClelland, 1987). The widely shared perception of founders as desiring autonomy could therefore raise commitment-related concerns, increasing questions about the high costs of

replacement, central to employment decisions (Leung, 2014). Indeed, ample research has documented that similar interruptions in paid employment, such as movement towards part-time work or flexible work arrangements, increase employers' perception that an individual is less committed and therefore less desirable as a candidate (Pedulla, 2016; Weishaar, 2018). And while entrepreneurship is substantively different from a leave, it may be interpreted as an equivalent signal of an applicant's willingness (or lack thereof) to invest in the hiring organization and generate a corresponding reduction in employer interest. Hence, even if employers accept independent-minded employees, concerns that a candidate may desire a return to entrepreneurship will help generate similar negative associations for ex-founders.

Overall, we therefore expect that commonly shared cultural beliefs about founders will elicit negative stereotypes about these individuals' fitness for and commitment to wage work upon re-entry. And that the net effect of these concerns will be to reduce assessments of job candidates who are ex-founders relative to similarly qualified applicants without founding experience. Hence, relative to a comparably trained employee in a traditional firm (i.e. non-founder), ex-founders will be liable to receive a discount at the recruitment stage and this discount will be driven by employers' beliefs that an ex-founder lacks fit and/or commitment to wage work.

H1a: *Ex-founders will receive lower employers' evaluations than non-founders.*

H1b: *Employers will be more likely to infer unfitness or lack of commitment to wage work for ex-founders than non-founders.*

Stereotype Conflict: The Evaluation of Female Ex-Founders

The core tenet of our theory is that employers rely on their stereotypical understanding of founder attributes and that, while desired or valued in entrepreneurship, these stereotypes might raise concerns in wage employment. Therefore, we further expect that the penalty will be mitigated in situations where these negative founder stereotypes are rendered weaker or less influential during evaluations.

Ample research suggests that even negative stereotypes can be counteracted in the presence of other, contradictory information. The psychological theory of stereotype contrast, for example, posits that group-level assumptions can be negated by counter-stereotypical information, and this interaction

between conflicting stereotypes can help reduce discrimination in the labor market (Blair 2002; Blair, Ma, and Lenton 2001; Dasgupta and Greenwald 2001; Power, Murphy, and Coover 1996). For example, in an experiment evaluating job candidates, Pedulla (2014) found that the labor-market chances of Black men were significantly better when the candidate was also gay. Although gay men have been shown to be the subjects of discrimination (Tilcsik, 2011), gay Black men benefited from a nullification of the dominant negative stereotypes. Specifically, the association of Black men as threatening was negated by contrasting stereotypes of gay men as effeminate and weak. More generally, employers may discount one negative stereotype if it is offset by a second contradictory one. Applied to the evaluations of founders, this implies that employers will be less influenced by negative founder stereotypes when the candidate's resume offers contrasting evidence in support of the applicant's propensity to commit and fit.

Applicant demographics, such as their gender or race, provide one of the most likely additional sources for stereotypes, and specifically for predictions regarding commitment and fit. Multiple studies have shown that, faced with the scarcity of objective evidence, employers invoke coarse demographic attributes to infer a candidate's preferences (Rivera and Tilcsik, 2016; Leung and Koppman, 2018). The effect of these actions often varies by the specifics of the context, as stereotypes can conjure both positive and negative associations during recruitment (Heilman, 1984; Kang et al., 2017; Pager and Quillian 2005; Pager and Shepherd 2008). However, gender and racial stereotypes prove so influential that even identical experiences in an applicant's history can acquire differing significance or meaning, depending on the applicant's demographic attributes (Rivera 2017; Rattan et al. 2017; Weisshaar 2018). Therefore, as a means of testing our proposed mechanism, we argue that entrepreneurship will prove less detrimental to job candidates whose demographic attributes contrast the founder stereotype and predict greater commitment and fit.

In particular, we identify two reasons that an applicant's gender is liable to bring to the fore contrary assumptions about their propensity for organizational commitment and fitness for wage work. First, long-standing evidence in social psychology suggests that, when they are subject to evaluation, women are commonly perceived as being more agreeable, more compliant, and more committed to their

employers (e.g., Fiske; 2002; Heilman, 1980; 1984 Ridgeway and Correll 2004). For example, in a study of the effect of gender and academic performance on hiring, Quadlin (2018) finds that employers exhibited less interest in determining the potential commitment or fitness for female than male candidates precisely because these attributes were associated with lower uncertainty for female candidates; conversely, participants questioned the potential fitness and likely commitment of male candidates. Further, employers view female candidates as less likely to leave for another employer because, given their disproportionate family obligations, women are less mobile and less willing to relocate (Rivera, 2017). While these gender-based stereotypes are not universally applied (Rosen and Jerdee, 1974; Terborg and Ilgen, 1975; Heilman, 1980), prior work has shown that their influence increases with their pertinence (Heilman, 1984). And, therefore, that an employer would be more likely to invoke a stereotype regarding female workers when doing so helps address specific questions around the candidate's suitability for the job (Heilman, 1984). By the same logic, the saliency of gender should increase when the candidate is an ex-founder, and concerns about a candidate's fit and commitment for the job become more significant, inclining employers to search for additional cues to confirm or disconfirm their initial assessments. As employers inquire for more information, stereotypical perceptions of women as more committed, more agreeable, and more dependent on others will likely come to the fore.³ More generally, negative stereotypes around ex-founders will be offset by gender stereotypes that portray women as more obedient, more committed and less rebellious than men.

Prevalent cultural beliefs about the entrepreneurial activity offer a second reason to anticipate that gender may moderate the negative assumptions attached to ex-founders. Entrepreneurship is highly gendered in that it is commonly associated with predominantly masculine traits (Thébaud, 2010) and believed to be incongruent with the feminine role (Kanze et al., 2018; Thébaud, 2010; Thébaud, 2015a). Indeed, gender changes the audience's interpretation and understanding of entrepreneurial efforts or intent

³ An alternative prediction would be that women are perceived as more committed to entrepreneurship because they face higher barriers to entry. While plausible, this argument, however, contradicts past studies showing that women are less likely than men to commit to founding new ventures, given the disproportionate time constraints and family responsibilities they face (Scott and Shu, 2017).

(Lee and Huang, 2018) so much so that women are even perceived as less entrepreneurial because they lack the masculine traits attributed to founders (Thébaud, 2010). Assessments of female entrepreneurs therefore often discount their experience or intent as less meaningful than that of comparable men. Consequently, even if employer's do retain questions regarding the candidate's unobserved attributes, the entrepreneurial traits commonly inferred from the experience of being a founder will function as a less relevant or less valued signal of those attributes. By contrast, because entrepreneurship is typically regarded as a masculine activity and men are more commonly seen as prototypical founders, any evidence of entrepreneurial experience will reinforce the preconceived notion of entrepreneurial traits when a job candidate is a man.

Taken together, we therefore expect that negative stereotypes about ex-founders will be mitigated for female candidates because (a) gender elicits positive stereotypes about organizational commitment and fit that will counteract negative perceptions; and (b) entrepreneurship is a male-typed activity and therefore employers will consider negative stereotypes about founding as less informative and less relevant when evaluating women. Female founders will therefore find a reprieve from one of the potential negative consequences of founding and their chances at the point of re-entering paid employment will be higher.

H2a: Male ex-founders will receive lower employers' evaluations than female ex-founders.

H2b: Employers will be less likely to infer unfitness or lack of commitment to wage work for female ex-founders than male ex-founders.

Data and Methodology

STUDY 1: Resume-based audit

To test our hypotheses, we first employ a resume-based audit approach, commonly used to assess baseline gender and racial biases net of any potential variation in the pool of applicants (Gaddis 2014; Pedulla 2016; Rivera and Tilcsik 2016; Tilcsik 2011). Consistent with common practice, realistic job applications are submitted for actual job postings, and the analyses aim to examine variation in the

requests for interviews.⁴ This is notably only a preliminary stage in the hiring process, but it is reflective of how candidates vary in the opportunities they encounter because submitting an application is a necessary condition when attempting to secure a job. Further, the use of interview requests rather than job offers reduces variability in the study design, eliminating any potential confounders that can arise at the interview stage and – due to their qualitative nature – are difficult to account for quantitatively.

In our study, we adopted a 2×2 (founder/employee × male/female) within-subject design in which we submitted two applications (one ex-founder, one non-founder) for each position that met our criteria (detailed below), varying randomly the gender of both applications. The virtue of this approach is that it is possible to test how founding a firm affects a prospective employer's interest in a given candidate, net of confounding factors. To limit the cost to employers, the outcome of interest is whether or not the employer contacts the candidate (by phone or email) for a follow-up interview. Although information about employers' propensity to hire an entrepreneur is not available, the probability of a callback provides initial evidence of variations in interest and is consistent with prior work that employed a resume-based audit approach (Correll et al. 2007; Pedulla 2016; Rivera and Tilcsik 2016). Significant to our specific question, it is possible with this approach to ascertain the influence of entrepreneurship on an employer's evaluation, and to assess whether this influence changes with the gender of the ex-founder.

Position Sampling

Between July 2017 and April 2018, we sent 1,223 applications for positions in two occupations and located in twelve major cities, selected from four different regions (Northeast, West, Midwest, South) of the United States, advertised on a leading job search website.⁵ To test the influence of local perceptions of entrepreneurship, we began by identifying twelve cities of roughly similar size and from similar regions but varying in their level of entrepreneurial activity. We then identified all full-time, non-entry

⁴ Automated responses and follow-up requests for online skill-based tests were coded separately and not included as an "interview request."

⁵ The process was conducted in accordance with an IRB approval granted by the second author's institution. As stated in our initial IRB application, the goal was 300 applications per condition, we exceeded this to account for a few positions that closed early.

level positions in marketing and HR (i.e. titles included “Manager”, “Director”, and “Coordinator”) within 25 miles of each city center; this included positions from five key economic sectors (e.g. IT, finance, health care).

The selection of the two occupations reflects our effort to balance multiple interests. First, we sought occupations in which entrepreneurship is a viable career option and therefore where both resumes are realistic representations of a potential career path. Second, we eliminated occupations that are male-typed or female-typed to alleviate a potential concern that any findings reflect gender role incongruence. Third, we sought occupations with a sufficient population of women to ensure that (1) our findings will apply to a significant group and that (2) any potential preference for women was not driven by underrepresentation. Given these constraints, we did not consider professions such as nursing—which has a high percentage of women, but offers few routes to founding—nor did we consider computer programming—which generates founders but employs few women. Finally, we needed occupations in which there was not such high levels of specialization that it would be impossible to create a resume that was acceptable to a range of positions. Applying our criteria led to the selection of HR and marketing, two of the most popular occupations for female college graduates, and also fields in which entrepreneurship is common.⁶ For example, over 18,000 HR or marketing-based startups received investments in 2016, and marketing is the second most common industry class (and HR the tenth) for firms in the Inc. 500 (Hathaway 2018). Interviews with experts in each of the occupational areas confirmed that there is mobility between these positions and that an applicant with our resume would be a viable candidate. By contrast, a pre-test determined that occupations like programmer or engineering were

⁶ In the United States, women are highly prevalent in HR, with 74 percent of workers being female (U.S. Department of Labor, 2016). Although this mitigates the concern that the occupation might be dominated by men, one might worry that men are underrepresented and therefore discounted. This concern, however, is alleviated because we also focus on Marketing, a field in which women and men are approximately equally represented, with female workers accounting for 45 percent of the workforce (U.S. Department of Labor, 2016).

so specific in their requested knowledge/background that it was difficult to construct a credible resume that would fit a wide variety of offerings without seeming artificial.⁷

Resume Construction

For every open position, we submitted two applications, and each resume was comprised of ten distinct elements (e.g. a first name, a last name, home address, first job, etc.). As explained below, we created two options for each city-occupation pair (e.g. the home address options were set by the employer's location) and assigned each value at random (within a city-occupation) at the time of submission—such that there was no correlation between any of the conditions and, for instance, attending a particular college or living at a particular address (Appendix A. Table 1). Only two variables: the first name and most recent job experience (i.e. position, firm, and activities) were fixed to a particular condition. We begin by explaining each of the varying elements.

First, to reduce the likelihood an employer discounted the second applicant, we created two sets of resumes using different fonts, styles, and arrangements of information (Figure 1 in Appendix A). Initial tests indicated no one aesthetic approach was preferred over the other. Second, given the evidence of a “home bias” among employers, we created “local” addresses that were specific to each city in our sample but not significantly different between cities. Specifically, each resume listed a distinct fictional address from adjacent target-city neighborhoods with similar demographic characteristics (e.g. SES, employment rate, education rate, racial composition). For example, the Philadelphia resumes indicated that the applicants lived in the 19130 and 19147 zip codes, which were not significantly different in racial composition or income, or even the average price for a one-bedroom apartment rental. Table 1 presents a breakdown of the four resume submissions by city, region, and occupation. Third, we selected two common last names from a list of the most common American surnames. Next, in a series of pre-tests, we identified pairs of large state universities that were similarly ranked by U.S. News, similarly selective,

⁷ Even seemingly entry-level positions like “Jr. Programmer” often required competency in a specific programming language and firms vary widely in which language they required (e.g. Python, Java, Apache, Ruby, etc.). A plausible IT audit would therefore require an applicant with competency in an unlikely variety of programming languages or multiple resumes each tailored to a given programming language.

and perceived as offering equally good training/educations.⁸ Each resume therefore listed the applicant as a graduate of a regional state university (e.g. Penn State or Ohio State) of similar reputation, and having majored in the applicable field (e.g., Management or Marketing, depending upon the available major). Finally, because individuals typically transition into entrepreneurship from established firms (Kacperczyk 2012; Sørensen and Sharkey 2014), we found pairs of equal-sized national firms headquartered in each of the respective cities and listed them as the applicant’s initial employers. The resumes indicated that the applicant had begun their career with a standard entry-level marketing/HR-track position at one of these national firms before receiving a small promotion (e.g. “Marketing Assistant” to “Marketing Analyst”). Because job-switching may be a signal in itself, both candidates left the national firm at the same time, but the control moved to a managerial position in a smaller firm while the treatment founded a firm. In sum, each employer received applications from two candidates who had worked for equal numbers of years, at similar positions, for a large employer; one then left to found their own firm, while the other left to accept a position with similar responsibilities in a smaller firm. Critically, these eight elements were allowed to vary such that there was no correlation between any of the options for a given city or occupation and any of the four conditions.

***INSERT TABLE 1 ***

Signaling Entrepreneurship

To signal entrepreneurship, we changed the most recent job experience and the email address of the applicant to indicate that they were founders or not. Most obviously, the ex-founder listed their most recent job experience as “Co-Founder” of a marketing or HR-related startup; while the non-founder (i.e. control) conditions listed their experience as a marketing/HR Manager at a mid-sized local firm. In addition, for ex-founders, the contact email address was listed as founder@..., while the non-founder listed a combination of their first/last name.⁹

⁸ In September 2016 we paid 600 participants on mTurk (95% HIT rate, US-based, 500+ prior HITS) to score the “quality of education” for four (out of 32) US public Universities. We then found pairs that did not vary significantly (and met our other criteria—similarly ranked and selective).

⁹ In case an employer investigated the startup further, we created a static webpage at the host site for the duration of the project indicating that the firm had recently been acquired by a competitor.

We made sure to indicate that the candidate was the co-founder of a new firm rather than self-employed. Growing research has recognized that self-employment differs profoundly from entrepreneurship (Budig 2006; Levine and Rubinstein 2017; Sørensen and Fassiottto 2011; Thébaud 2015a), as it involves “initiating a sole proprietorship to sell one’s own service or products,” (Carroll and Mosakowski 1987). We thus specified on the resume that (a) the applicant’s most recent (third) position title was co-founder; and that (b) the startup “employed 10-12 people.” Both of these signals are inconsistent with self-employment. To reduce the likelihood that employers interpret entrepreneurship as developing unique skills, we described the activities and responsibilities of both the control and treatment positions as equivalent. For example, both applicants worked directly with clients and oversaw the development of the company’s strategy. To ensure that the only significant difference between two candidates was their entrepreneurial experience—and not an accidentally suggested skill or quality difference—we had the resumes evaluated first by professional HR managers (July 2019) and later by 600 mTurk workers (August, 2020). In each case, the respondents were shown a single resume and then asked to rate the skills and quality displayed on a 5pt scale. Analysis of these evaluations confirm that there were no significant differences in the ability or quality of the founder and non-founder resume.

Signaling Gender

To assess whether our effects varied by gender, we created male and female resumes, each with a treatment/control condition indicating that the person had/had not founded their own firm. To signal gender, we altered the first name of the applicant and changed the listed “interests” to activities determined in a pre-test to reflect similar SES but be more associated with women than men (e.g. baseball to softball). The names were selected via a pre-test in which 500 respondents were shown a random set of eight names and asked to indicate the gender, ethnicity, and level of education for a person with the given name (Gaddis 2017).

Application Process

For each position, we submitted a non-founder (i.e., control) resume and a founder (i.e., treatment) resume. The resumes were sent approximately 1-2 days apart to avoid raising suspicion, and the order of submission was randomized so that there was an equal chance of sending the founder first or second. In practice, this meant that any position could have received any combination of genders but always received one with an entrepreneurial career stint and one without it. When we found a posting that met our criteria, we randomly selected either the "A" or "B" style of resume and then one of the four potential conditions (i.e., female entrepreneur, female non-entrepreneur, male entrepreneur, male non-entrepreneur). The paired resume then assumed the opposite values (e.g. "B" if "A" were selected; founder if non-founder) with the gender always assigned at random such that it was possible to have two female, two male, or one male and one female applicant for each position. In approximately 5 percent of instances, the positions were closed before the second resume was submitted but these closings occurred evenly to the application types and did not reflect any systematic differences in the analyses. For robustness, we excluded these applications from our analyses and found similar results (available upon request). A typical HR application pair is displayed in Appendix A, Figure 1 (an example of the male and female marketing resumes is provided in Appendix A, Figure 4 & 5).

Results

Our first hypothesis proposed that employers would be less likely to interview an ex-entrepreneur than a similarly qualified candidate who had worked for a traditional organization (i.e. not a start-up). The results of the audit lend support to this hypothesis. At the broadest level, when pooling both female and male applicants (Appendix A, Figure 2), a mean comparison test shows that ex-founders were significantly less likely to be called in for an interview (mean=0.105, S.D.=0.306) than non-founders (mean=0.167; $p=0.002$). The results are not sensitive to approach and both ANOVA and OLS report identical evidence of a founding penalty.

Figure 1 depicts these comparisons separately for men and women. Given the gendered nature of entrepreneurship, we begin by considering the experience of male candidates. As hypothesized, male ex-founders were significantly less likely to be called in for an interview (mean=0.072, S.D.=0.259) than

male non-founders (mean=0.160, $p=0.001$). These results are therefore consistent with our prediction that the act of founding is associated with negative employers' perceptions, thus reducing the number of opportunities at the point of re-entry into paid employment.

Our second hypothesis (H2a) suggested that this baseline penalty would not apply to job candidates who trigger a different interpretation of entrepreneurship, such as female founders. The results presented in Figure 1 support this proposition, showing that women with an entrepreneurial experience were significantly more likely to be called in for an interview (mean=0.138, S.D.=0.345) than male ex-founders (mean=0.072, S.D.=0.259; $p=0.008$).¹⁰ Further, this finding does not vary by occupation as in both HR and marketing male founders were equally (6 percent in each case) and significantly less likely to receive interviews than female founders. In addition, while female ex-founders were called in for an interview (mean=0.138, SD=0.345) at a lower rate than female non-entrepreneurs (mean=0.174, SD=0.380; $p=0.22$), the difference was not statistically significant. In combination, these results provide some initial evidence that employers exhibit bias against ex-founders but also that women were not subject to the same entrepreneurial penalty that men faced.

INSERT FIGURE 1

While the resumes were submitted in pairs and at random to achieve balance across the conditions, it remains possible that the above findings reflect consistent differences in the positions or between the firms. To mitigate potential confounders, we performed a secondary analysis that estimated the likelihood of a given resume receiving a request for an interview controlling for the observed differences in the positions and applications. In a regression framework, it is possible to account for variation in aspects of the position or other external factors. First, we account for differences between the job postings to limit the possibility that, despite our randomization, such factors could still be unequally balanced across the treatment and control groups and thus affect the hiring outcome. The observed variables, and a brief definition, are summarized in Table 2. Specifically, we measured the number of

¹⁰ Because we find no difference in callback rates between female entrepreneurs and non-entrepreneurs, our results are unlikely to reflect employers' motivation to increase women's representation in the applicant pool.

years of specialized experience, either in marketing or HR, requested by the employer to account for a concern that an entrepreneurship stint might signal that job candidates have had a shorter experience in the labor market. We next constructed dummy variables to measure whether the position title was mid-level, or senior—as specified by the job posting site—to mitigate the concern that an employer-individual seniority match was driving our results. Eighty-nine percent of our jobs were classified as mid-level. We also included a binary covariate to measure whether the employer required applicants to complete EEOC surveys in order to control for the possibility that our estimates regarding gender could be driven by diversity initiatives.

We also collected the full text of each job posting to control for differences in the specified preferences of each employer. For instance, we hand-coded (binary) whether or not the posting mentioned (1) an interest in "entrepreneurial" people and whether it listed (2) "leadership" or (3) "autonomy/independence." These covariates were included in our models to mitigate a concern that job-person fit might influence the employer's propensity to discount ex-founders. Similarly, we extracted from the job texts the information about whether employers sought a "generalist" by coding "1" if the text mentioned "generalist" skillset, and "0" otherwise. Generalism may again influence the call-back rate to the extent that employers may perceive founders to have more general skills. Next, we coded "1" if jobs disclosed salary or salary bands, and "0" otherwise, because such jobs may, on average, be higher-paid and this may, in turn, impose a higher selection bar on job candidates, leading to employer discounting of ex-entrepreneurs.

We also accounted for firm age and size. Because firm size was highly skewed, we log-transformed this measure before including it in our models. In cases where information on firm size and firm age was missing, we hand-collected these data from the company websites. In a few cases (about 10 percent of our observations) where the information was still missing, we gathered data from the National Establishment Time-Series (NETS) database, which provides annual information on public and private

companies in the United States, starting from 1989.¹¹ To account for regional differences, we created an ordinal measure to account for the amount of entrepreneurship within a city, based on the level of VC funding (PWC, 2018). Cities were scored as a “3” if the per capita investment exceeded \$1,000 (top 5 US city), “2” if it exceeded \$100 (top 20), “1” if it exceeded \$50, and “0” for all cities below \$50.

In addition, we created a dummy variable to indicate whether the resume was the first or second we submitted to a given firm; counted the number of days since the job was first posted; and, controlled for the fiscal quarter. Finally, because we applied for positions in two different occupations that may vary in the overall rates of interviewing, we included an indicator variable for whether the position was in HR (1) or marketing (0). We also included dummies for each of the five market sectors (i.e., Utilities, Services, Finance, Health, and IT).

The descriptive statistics and pairwise correlation table for these variables are presented in Table 3. In Table 4, Models 1-2 present naïve approaches with no controls or fixed effects to show that our results are not unique to one specification. Models 3-6 include indicator variables for city and fiscal quarter; as well as position and market sectors. Standard errors are clustered by region and by job advertisement because intra-class correlation may arise within regions and job advertisements (our results are robust when we cluster standard errors less conservatively, by city or by region alone).

INSERT TABLE 2, TABLE 3 AND TABLE 4

In Model 1 of Table 4, we replicate our test of the first hypothesis—an entrepreneurial penalty—estimating the results with a logit analysis given the binary dependent variable. Model 1 represents the univariate regression model, and yields results consistent with H1, showing that employers are less likely to interview ex-founders ($\beta = -0.54; 0.19$). In Model 2, we introduce an interaction between female and entrepreneurial applicant to determine whether the effect of the entrepreneurial penalty is moderated by

¹¹ Our results (available upon request) are also robust to the exclusion of these missing observations, as we obtain quantitatively and qualitatively similar results.

the gender of the applicant. Our results confirm this hypothesis, showing (Figure 3, Appendix A) that male ex-founders incurred greater penalty than female ex-founders at the pivotal point of hire ($p < 0.01$). In Model 3, we include in our models fixed-effects estimators (i.e., for city, fiscal quarter, position level, and market sector). But even after the inclusion of these effects, our results continue to hold, with the coefficient of the interaction term decreasing by 4.1 percent. In Model 4, we add additional controls for job-posting characteristics, application timing, and employer characteristics. As can be seen, this does not reduce the significance and economic magnitude of the *Ex-Founder* penalty, most likely because these additional covariates are balanced across the treatment and control groups.¹² A marginal effects calculation indicates that having entrepreneurial experience is associated with a 61 percent increase in the odds of receiving a callback, relative to having no past entrepreneurial experience [$=\exp(-0.926)-1$, $p < 0.001$]. In addition, the interaction term between our treatment and gender is statistically significant at 1 percent level. However, as noted by Ai and Norton (2003) as well as Hoetker (2007), interpreting interaction effects in nonlinear models must be done with caution because it requires taking into account not only the coefficient on the interaction variable and those on the interacted variables but the values of all other variables in the model. Hence, to further verify that entrepreneurial experience and gender were indeed significant, the algorithm of Norton et al. (2004) was used.¹³ When applied to Model 4, this algorithm yields an average interaction effect of 4.4 percent. The effect is significant at the 5 percent level, with a mean z-statistic of 1.96. These findings support the significance of the interaction term, suggesting that the negative effect of entrepreneurial experience on the callback rate is indeed reduced for female job applicants. Finally, in Model 5, we re-estimate our analyses from Model 4 but estimate the

¹² In additional analyses (unreported), we re-estimated our baseline specifications with a fixed-effect estimator. Although doing so results in a significant loss of observations (82%), we are still able to recover our results. The treatment effect continues being negative and statistically significant ($p < 0.01$). Similarly, we find a positive and statistically significant effect ($p < 0.01$) for the interaction term between treatment and gender. Finally, we re-estimated our baseline specifications with a random-effect estimator. Again, we were able to recover our results, with the treatment being negative and highly statistically significant and the interaction term with gender being positive and statistically significant at conventional levels ($p < 0.1$).

¹³ This method computes the correct marginal effect of a change in two interacted variables by calculating the cross-partial derivatives required to evaluate a single two-way interaction effect for each observation.

probit model as an alternative specification. As can be seen, we are able to recover our results. In addition, when applying the algorithm of Norton et al. (2004), we obtain the average interaction effect of 4.81%. The effect is again statistically significant at 10 percent, with a mean z-statistic of 1.72 and the values ranging from 0.618 to 5.50. Together the two sets of analyses support our first two hypotheses and provide evidence that employers penalize prospective HR or marketing employees for pursuing entrepreneurship, but also that this penalty is less pronounced for women than for men. These initial results do not yield any insights into the motivation for the penalty, but they provide evidence that women pursuing careers in HR and marketing are less likely to be disadvantaged than men due to an entrepreneurial career history. Below, we consider some alternative explanations for the baseline penalty and the gender difference, and then introduce a second study to further demonstrate our theory.

Supplemental Analyses

Diversity Preference. One concern is that, if employers desire a diverse applicant pool, they may give preferential treatment to female applicants when selecting interviewees. To mitigate this concern, we intentionally selected occupations in which female candidates are well represented and therefore it is less difficult for employers to generate a gender-balanced pool. In addition, if employers generally favor female candidates, then we should see a baseline preference for female candidates. However, a one-way ANOVA shows no statistically significant difference in the response rates for those two sets of applications ($\beta = 0.24$; $p = 0.626$). Finally, if concerns around representativeness motivated the selection of female candidates, then we might expect that, when firms signalled EEOC compliance or required EEOC questionnaires, they would be more favorable to women than a typical firm. However, Model 1 of Table 2 in Appendix A shows that the interaction term between the ex-founder status and the EEOC requirement is not statistically significant at the conventional levels.

Employer and Job Heterogeneity. We theorize that employers will penalize founders because founding suggests unwanted personal traits—an inability to fit or commit. If our theory is plausible, then we should observe that the baseline penalty will be mitigated for positions which value entrepreneurial experience or traits. Similarly, we would expect that the penalty for a founder's experience might be

amplified for larger and older firms, based on the accumulated evidence that these kinds of organizations are more rigid and more bureaucratic (Gompers, Lerner, and Scharfstein, 2005; Sørensen, 2007; Elfenbein, Hamilton, and Zenger, 2010). Appendix A Table 2 Models 2-3 show the estimates for these predictions. Consistent with our expectation, the interaction term between positions that mention entrepreneurial traits and ex-founder status is positive and statistically significant at conventional levels ($p < 0.01$), indicating that the negative effect of founder status is dampened for positions which value entrepreneurial experience more. When applying the Norton et al. (2004) algorithm, the average interaction term is 3.22 percent and is significant at the 1 percent level, with a mean z -statistic of 3.08. These findings support the significance of the interaction term, suggesting that penalty for entrepreneurial experience is mitigated for jobs which seek and value a founder's experience. Similarly, the interaction with firm age is negative and statistically significant ($p < 0.05$), indicating that, as expected, older organizations are more likely to penalize former founders. When applying the Norton et al. (2004) algorithm, the average interaction term is 1.1 percent and it is significant at the 1 percent level for most observations, with a mean z -statistic of -2.10. Hence, the result is robust to the correction. Finally, these results (unreported) are robust when, instead of using a continuous variable for firm age, we construct a dummy variable equal to "1" when firm age is above the median value in the sample. In addition, entrepreneurial experience is statistically significant at conventional levels when interacted with the request for leadership skills ($p < 0.1$). When applying the Norton et al. (2004) algorithm, the average interaction term is 5.4 percent and it is significant at the 5 percent level for most observations, with a mean z -statistic of 2.30. In sum, these analyses offer at least partial support to the notion that employers penalize job candidates for their entrepreneurial experience and that this penalty is mitigated when job and employer characteristics increase the value of entrepreneurial experience.¹⁴

Over-qualification. Another alternative explanation might be that founders receive a penalty at the point of hire because they are perceived as overqualified for the focal positions. It could further be

¹⁴ This heterogeneity also implies that employers are not suspicious of the founder resume, but that they are attentive to it and respond to the information therein.

that this penalty is mitigated for women because employers are less likely to evaluate overqualified female applicants negatively (Galperin et al. 2019). Although plausible, this explanation is unlikely given that our experiment relied on mid-level and senior-level positions, which might be a more suitable match for job candidates with entrepreneurial experience. Moreover, our findings indicate that the penalties for male founders are mitigated for positions and firms less likely to discount entrepreneurial traits, whereas overqualification would imply a uniform penalty across such jobs. Nevertheless, to additionally address this possibility, we replicated our models within the subsample of managerial jobs. We identified managerial jobs as positioned with the mention of “Manager” in the job title. If the initial results reflect an overqualification penalty then we would expect the baseline penalty to no longer hold. Although our sample is reduced by 70 percent, we recover our results: Model 3 of Table 2 in Appendix A shows that the treatment continues being negatively associated with the outcome ($p < 0.01$), and this effect is dampened for female ex-founders ($p < 0.01$).

Gender and Skill. Our findings indicate that the penalty for ex-founders is mitigated when candidates are women. We attribute this finding to the fact that stereotypes associated with gender – that women are less entrepreneurial overall and that they are more committed and fit for wage work – weaken the negative stereotypes about founders. An important concern, however, might be that recruiters perceive ex-founders as less capable or less skilled job candidates, but that women may be exempt from such presumption, because recruiters give them credit for having overcome the barriers, which are thought to place women at a systematic disadvantage in entrepreneurship. While plausible, this alternative explanation is unlikely because we do not find a premium for female founders – which would be expected if recruiters perceive women who started their own ventures as more capable. Nevertheless, we take additional steps to mitigate this concern by leveraging cross-sectional heterogeneity in job postings. If employers infer greater skill for female ex-founders, then chances of these candidates should additionally increase for jobs with a higher selection bar. One factor which increases the selection bar is the years of experience required by employers. However, Model 4 of Table 2 in Appendix A shows that the triple interaction term between the ex-founder status, gender and experience requested by employers is not

statistically significant at the conventional levels and continues to be so when we apply the Norton et al. (2004) algorithm. Together with other evidence, this non-finding provides some reassurance that differences in skill are unlikely to account for the diminished gender penalty we document.

STUDY 2: Survey Experiment of Hiring Managers

We next turn to a survey experiment to assess the evaluators' perceptions and test our H1b and H2b more directly. Our theory suggests that employers will apply lower penalty to female than to male founders in part because negative founder stereotypes will be rendered weaker in presence of other stereotypes.

Survey Design

To explore these perceptions, we recruited 425 marketing managers through a Qualtrics panel to take part in a survey experiment, following the well-established procedure in prior studies (Galperin et al. 2019; Rivera, 2016; Parigi, Santana, and Cook, 2017). All participants had extensive recruiting experience hiring for marketing positions in the United States.¹⁵ A majority of managers were female (53 percent), had more than five years of job experience (89 percent), and a bachelor (93 percent) or graduate (10 percent) degrees. The participants were drawn from across the U.S. and across a range of industries, such as manufacturing, advertising, healthcare, software development, and consulting.

The survey experiment used a 2×2 between-subject design, presenting each participant with one of four possible resumes, all drawn from our sample of Minneapolis marketing conditions (see Appendix A, Figures 4 & 5). Participants were asked to use the resume to evaluate the applicant's fit and commitment and were told that their inputs would help make better recommendations.¹⁶ After agreeing to evaluate the resume, each participant was randomly shown one of the four potential conditions (*male employee, male founder, female employee, female founder*), and then asked to evaluate the candidate on a number of dimensions (detailed below) using a 7pt. Likert scale. After offering an overall

¹⁵ We did not have direct contact with the participants but instead relied on the Qualtrics firm to recruit participants who were U.S.-based hiring managers.

¹⁶ Each participant was told that reading the resumes and answering all the questions will take about 10 minutes of their time. The specific instructions are reproduced in Appendix B, Figure 1.

recommendation (and before rating questions), the recruiters were asked to explain their rating in an open-ended question. Finally, each participant was asked an attention check and the mechanism check. We subsequently removed from the final sample any participants who did not complete the survey or participants who failed to pass the attention check.¹⁷ This led to a final sample of 413 individuals.

Dependent Variable. The principal outcome we modelled was the respondents' *Willingness to recommend the applicant (as a hire) to a prospective employer*. Specifically, each respondent was asked, "Would you recommend this job candidate to a future employer?" Replies were on a scale ranging from "1" (definitely will not) to "7" (definitely will). To gather qualitative data (as discussed later), the initial rating was followed by an open-ended question: "Please explain your recommendation." This question was asked before we turned to closed questions in order to avoid biasing our respondents.¹⁸

Independent variable. The main predictor variable was *Ex-Founder*, coded as "1" if the resume belonged to an ex-founder, and "0" otherwise.

Moderating variable. The main moderating variable is *Female*, coded as "1" if the resume belonged to a female job candidate, and "0" otherwise.

Mediating variables. To assess whether inferences about fit generate a penalty for ex-founders, we directly asked participants to assess whether they believed the job candidate "was a good fit for a more traditional/hierarchical organization," coding their replies on a 7 pt. scale ranging from "1" (definitely no) to "7" (definitely yes). Complementing this broad measure with tests of specific fit dimensions, we asked employers to assess the likelihood that the job candidate will "quit their next job," (7 pt. scale; "definitely will not" to "definitely will") consistent with prior research which has conceptualized commitment as a continuous relationship with an employer (e.g., Pedulla 2020). Second, because ex-founders might be seen as placing high value on autonomy and independence (Wasserman, 2006; Hamilton, 2000; Benz, 2009; Hurst and Pugsley, 2011), we measured the perception that a job

¹⁷ As the manipulation check, we asked the participants to recall whether a job candidate was an ex-founder and eliminated participants with incorrect responses. This reduced the sample size by 2.8 percent.

¹⁸ Qualitative evidence (available upon request) confirms the quantitative analyses, with respondents viewing ex-founders as potentially less fit and de-committed employees.

candidate was “*easy to manage*,” (7 pt. scale, “definitely not” to “definitely yes”). Finally, given that ex-founders might be perceived as unlikely to be fit for teamwork, given their desire for independence (e.g., Moskowitz and Vissing-Jørgensen, 2002), we measured the perception that a job candidate was “*a team player*,” (7 pt. scale, “definitely not” to “definitely yes”).

Empirical Specification

All models include controls for the respondents’ characteristics, such as gender (“2” if the participant was female, and “1” otherwise), age (continuous variable), and years of work experience (dummies for experience categories, including “less than 5 years;” “5 to 10 years;” “10-20 years;” and “over 20 years”). Our models also included dummies for the participant’s location at the region level (“Northeast”; “Midwest”; “South” and “West”), as well as dummies for their educational level (i.e., “Bachelor,” “Masters”; and “PhD”).

Main Results

Descriptive statistics and pairwise correlations are presented in Appendix B, Table 1 and 2. We begin by analyzing differences in means between our treatment and control conditions. Consistent with the audit results, participants gave a significantly stronger ($t=2.05$; $p=0.04$) interview recommendation to non-founders (mean = 5.73) than to ex-founders (mean = 5.40), and female ex-founders received a significantly higher recommendation ($t=4.33$; $p=0.01$; mean=5.89) than male ex-founders (mean=4.86). Table 5 repeats this analysis with a regression framework and displays OLS models predicting interview recommendations as a function of treatment.¹⁹ Model 1 replicates the estimates from the audit study on the full sample of job candidates: the estimates are recovered, with the coefficient of *Ex-Founder* being negative and statistically significant at the 5 percent level. Model 2 re-estimates this baseline specification but adds the interaction between *Ex-Founder* and *Female*: the coefficient of *Ex-Founder* is statistically different for female and male job seekers, as indicated by the positive and statistically significant interaction term between *Ex-Founder* and *Female Candidate* ($p<0.05$). The magnitude of this effect is

¹⁹ For robustness, we estimated ordered logit models and found similar quantitative and qualitative results.

striking: Among male candidates, the probability of being recommended for an interview decreases by .765 points on a 7-point Likert scale.”

INSERT TABLE 5

Quantitative Analyses: Mechanisms

Perceptions of Fit. We next turn to the causal mediation analyses to probe the underlying mechanisms described in H1b and H2b. H1b predicts that ex-founders might face a penalty based on fit in wage employment while H2b predicts that this penalty will be mitigated for female founders. To test these predictions, we conduct a moderated-mediation analysis, according to which gender differences in obtaining recruiters’ recommendation are affected by the intervening variable of perceived fit. We follow a common practice (Rosnow and Rosenthal, 1989; Preacher, Rucker, and Hayes, 2007) and implement a counterfactual-based approach to mediation, which allows us to decompose the total effect into a direct and indirect effect even in the presence of exposure-mediator interaction. We begin with the joint model for men and women (Table 6 Panel A), followed by a subsequent analysis of “separate groups” to confirm that the proposed theoretical path differs by gender, as stated in H2b (Table 6 Panel B and Panel C).²⁰ Table 6 Panel A examines the joined OLS model, estimating the relationship between the treatment (i.e., being an ex-founder) and the mediator (i.e., employers’ perceptions of job candidate’s fit).²¹ Model 1 in Panel A establishes the association between *Ex-Founder* and hiring recommendation, whereas Model 2 shows significant differences in the perceived fit of female and male founders, as indicated by the positive and statistically significant coefficient of *Female Candidate* $\times$ *Ex-Founder* ($p < 0.05$).

To examine whether the indirect effect of *Ex-Founder* $\times$ *Female Candidate* via perceived fit on the hiring recommendation is significantly different from zero, we bootstrap standard errors with a sample of 5,000 replications (Preacher and Hayes, 2008). The indirect effect is statistically significant ($p = 0.033$),

²⁰ To conduct the causal mediation analysis, we implement the counterfactual approach based on the SPSS macros, which allows for the estimation of direct and indirect effects in presence of interactions (Valeri, 2012). This counterfactual approach accounts for the effect of potential confounding factors.

²¹ We used the STATA Paramed command (Emsley et al. 2014), which implements the same analyses as the SAS and SPSS macros.

providing support to our hypothesized model in the joint analysis of men and women. The significant indirect effect found in the moderated-mediation analysis for the full sample implies that the impact of negative stereotypes differs for men and women. To further test this, we conducted separate group analyses for women and for men (Panel B of Table 6) and find that, as expected, a history of founding was not significantly associated with negative stereotypes when the candidate was a woman ($p > 0.1$ in Model 2), even though fitness was significantly associated with the hiring recommendation ($p < 0.01$ in Model 3). We also tested for the indirect effect within the female subsample in another bootstrap sample of 5,000 and found that the indirect effect was not statistically significant ($p = 0.427$). By contrast, when we conducted a similar analysis on the subsample of men, we found that the indirect effect—via perceived fitness—of entrepreneurial experience on employer recommendation was statistically significant ($p < 0.01$). At the same time, the direct effect is only borderline significant: a p -value equals to 0.082. Finally, Model 3 in Panel C shows that controlling for the perceived fit reduces the economic magnitude of *Ex-Founder* treatment by 46 percent, while the statistical significance decreases to 10 percent. Figure 1 in Appendix B shows the causal paths in the moderated-mediation analysis for the joint sample of women and men. Figures 2-3 in Appendix B show the causal paths for men (Figure 2) and women (Figure 3) separately.

Overall, these analyses provide support for H1b, suggesting that the negative link between the history of founding and hiring recommendation is mediated by the recruiters' perception of fitness. We also find support for H2b, as our results indicate that the history of founding is more likely to trigger the perception of being ill-fit for wage work among men than women. Finally, these differential perceptions lead recruiters to discount male ex-founders more than female ex-founders.

INSERT TABLE 6

Dimensions of Fit. Our analyses establish that perceptions of being unfit for wage work drive the penalty that entrepreneurs receive upon their return to wage work. In additional analyses, we probe deeper into such perceptions, identifying specific attributes that employers associate with entrepreneurial experience. Table 7 re-estimates the baseline analyses from Table 6 but considers the specific negative

stereotypes recruiters may hold, including (a) the likelihood to quit in Models (1)-(2); (b) being easy to manage in Models (3)-(4); and (c) being a team player in Models (5)-(6). These additional analyses are consistent with our expectations. Models (1)-(2) establish that employers perceive ex-founders as being “more likely to quit” than non-founders ($p < 0.001$), and that these differences are mitigated for female ex-founders, as indicated by the negative and statistically significant interaction term between *Female Candidate* and *Ex-Founder* ($p < 0.05$). The moderated-mediation model, which involves examining the indirect effect of *Ex-Founder* $\times$ *Female Candidate* via quitting on the hiring recommendation was significantly different from zero: the bootstrap sample of 5,000 replications indicated that the indirect effect is statistically significant ($p = 0.080$), providing support for our hypothesized model in the joint analysis of men and women. A separate-group analysis (available upon request) showed that ex-founder status was positively associated with the perception of quitting in the female subsample ($p < 0.1$) as well as the male subsample ($p < 0.01$). A bootstrap sample of 5,000 within female and male subsamples further revealed that the indirect effect was not statistically significant for women ($p = 0.012$) but was so for men ($p = 0.016$). Together, these results show that entrepreneurial experience is more likely to trigger the perception of a job candidate being likely to quit but that this effect is weaker when a candidate is female.

Second, Models (3)-(4) of Table 7 report that recruiters perceived founders as being less “easy to manage” than non-founders ($p < 0.001$), and that these differences were again mitigated for female ex-founders, as indicated by the positive and statistically significant interaction term between *Female Candidate* and *Ex-Founder* ($p < 0.05$). The moderated-mediation model was also significantly different from zero: the bootstrap sample of 5,000 replications indicated that the indirect effect is statistically significant ($p = 0.089$). A separate-group analysis (available upon request) lends additional support for our predictions. As can be seen, the ex-founder’s status was negatively associated with the perception of “easy to manage” for both female and male ex-founders. A bootstrap sample of 5,000 within female and male subsamples further revealed that the indirect effect was statistically significant for men ($p = 0.001$) but not women ($p = 0.108$). Thus, recruiters perceive ex-founders as “uneasy to manage” but more so in the case of men than women.

Finally, Models (5)-(6) of Table 7 establish that employers' perceive ex-founders as less likely to be a "team player" than non-founders ($p < 0.001$) and that these differences are mitigated for female ex-founders, as indicated by the positive and statistically significant interaction term between *Female Candidate* and *Ex-Founder* ($p < 0.01$). The moderated-mediation model was also significantly different from zero: the bootstrap sample of 5,000 replications indicated that the indirect effect is statistically significant ($p = 0.070$), providing support for our hypothesized model in the joint analysis of men and women. In a separate-group analysis (available upon request), we further find that being an ex-founder was positively associated with the perception of being a "team player" in the female subsample ($p < 0.1$) but negatively in a male subsample. A bootstrap sample of 5,000 within female and male subsamples further revealed that this indirect effect was not statistically significant for women ($p = 0.917$) but that it was significant for men ($p = 0.002$). Together, these results provide additional support for H1b and H2b. They establish that the specific negative stereotypes about commitment and fit (i.e., being more likely to quit, being easy to manage, and being a team player) mediate the relationship between the founding history and recommendation to hire (H1b), but more so for men than women (H2b).

INSERT TABLE 7

Masculine Stereotypes of Entrepreneurship. The second proposed reason for why negative founder stereotypes become weaker when job candidates are women pertains to the notion that entrepreneurship is a male-typed activity and that women are seen as less fit for entrepreneurship, in general. Similar to investors or other stakeholders who hold a negative bias against women (Thebuad, 2010; Guzman and Kacperczyk, 2018), recruiters may also exhibit such bias, seeing female founders as less legitimate or less qualified entrepreneurs. We probe this explanation deeper by asking the respondents whether "*The candidate was a good fit for their last job (i.e., the job preceding the current job search).*" Replies were on a scale ranging from "1" (definitely was not) to "7" (definitely was).

Table 8 reports estimates for differences in the perception of fit into entrepreneurship by job candidate's gender. Model 1 shows that the interaction between *Ex-Founder* and *Female Candidate* is negative and statistically significant at the 5 percent level. Models 2 and 3 further estimate these

associations for the subsample of women (Model 2) and men (Model 3). As can be seen, within the subsample of female candidates, the coefficient of *Ex-Founder* is negative and statistically significant at the 5 percent level, whereas, within the subsample of male candidates, our treatment is not statistically significant. Overall, these results provide evidence consistent with the notion that female ex-founders are more likely than male ex-founders to be perceived as unfit for entrepreneurship.

INSERT TABLE 8

Alternative Explanations

Perceived Skill Difference. The experimental study provides direct evidence that perceptions of commitment and fit mediate the relationship between founding and hiring recommendations more for men than women. An alternative explanation, however, may be that recruiters perceive female founders as more capable or better skilled. Although we found no evidence consistent with claim in the audit study, we additionally probed this mechanism using the experimental data. Specifically, we asked the respondents whether they believed that “*the job candidate was highly skilled.*” Replies were on a scale ranging from “1” (definitely not) to “7” (definitely yes). To conserve space, these results are reported in Appendix A, Table 3. As can be seen in Model 1, perceived competency is not significantly associated with the *Ex-Founder* treatment; similarly, no statistical differences were detected in perceived skills across female and male ex-founders ($p > 0.1$), even though perceived skills are positively associated with the recommendation to hire, as shown in Model 2 ($p < 0.01$). Together, these results mitigate a concern that our results are driven by differences in perceived skills.

Perceived Failure. A related alternative explanation is that our results simply reflect an inference of failure. It may further be that women are excused from a failure-based penalty if audiences perceive them as facing greater obstacles as founders. We investigate this possibility, by examining whether recruiters inferred failure from the stint of entrepreneurship. Specifically, we asked the respondents whether they believed that “*the job candidate failed in their last job.*” Replies were on a scale ranging from “1” (definitely not) to “7” (definitely yes). As can be seen in Model 3, perceived failure in the last job is positively associated with the former ex-founder status, suggesting that recruiters tend to infer

failure from an entrepreneurial spell more than from a regular job spell. However, our results indicate that gender is unlikely to moderate this perception, as indicated by the statistically insignificant coefficient of the interaction term in Model 3. Furthermore, in Model 4, the coefficient of the *Failure* variable is not statistically significant, suggesting that recruiters, despite perceiving ex-founders as having failed, are less likely to penalize them for failure.

Salience of Ex-Founder Status. It is further possible that the gendering of the penalty does not reflect variation in the meaning of entrepreneurship (as we propose) but simply variation in salience. For example, it might be that employers are more likely to fail to notice entrepreneurship when the founders are female. By contrast, an entrepreneurial stint might be more salient or easier to recall for men. Thus, prospective employers may be equally disposed to penalizing male and female founders, but they do not register the signal as easily for women. We address this concern in two ways. First, all of our analyses exclude respondents who failed the manipulation check, or those unable to correctly recall whether the job candidate was a founder or simply an employee. Second, to probe the possibility that salience of entrepreneurship might vary between female and male candidates, we additionally asked the respondents the following question: “*What are the top three things you remember about the job candidate?*” We then ask the participants to select three items from the following list: “Their college,” “Their degree,” “Their location,” “Their name,” “Their gender,” “They were a founder,” “They changed jobs,” “They were promoted,” “They barely changed any jobs.” Based on this list, we constructed a dummy variable equal to “1” if a respondent selected “They were a founder” as one of the three choices. Our analyses reveal two important conclusions. First, the majority of respondents (fifty four percent) who evaluated an entrepreneurial resume included entrepreneurial experience among the top three attributes they recalled about the job candidate, suggesting that entrepreneurial experience was highly salient to respondents. Second, respondents were not more likely to recall entrepreneurial experience for male than female entrepreneurs. Model 7 recovers this finding in the logit regression specification, with the outcome being a binary variable equal to “1” if ex-founder is mentioned as one of the top recalled job candidate’s

attributes. As can be seen the female dummy is positive and statistically significant at the 10 percent level. This finding indicates that founding is salient for both female and male candidates.

Necessity Entrepreneurship/Self-Employment. Another concern might be that our main effect reflects inferences about necessity entrepreneurship, whereby recruiters perceive ex-founders as workers with interrupted employment (e.g., Astebro et al., 2014; Conti, Kacperczyk and Valentini, 2019). If so, then the penalty we document may not be tied to the pursuit of entrepreneurship specifically but instead reflect a more common aversion to unemployed or part-time candidates. However, there are theoretical and empirical reasons to believe that entrepreneurship is recognized as a distinct and unique activity. First, entry out of necessity typically involves limited operations, often without formal employees or within home-based locations (Aldrich 2005; Levine and Rubinstein 2017; Ruef 2010; Steinmetz and Wright 1989). To ensure that our resumes did not signal necessity entrepreneurship, our treatment specified that individuals (a) started business with co-founders; and (b) the firm employed ten to twelve people. As an additional test, the respondents were asked to choose the most likely reason the individual left their prior job but only four participants (0.02 percent) presumed that the ex-founder was fired, mitigating the claim that entrepreneurship was seen as equivalent to unemployment.

Discussion

Our study investigates the effect of pursuing entrepreneurship on the future employment of male and female HR or marketing professionals. Drawing jointly from demand-side approaches to hiring (e.g., Bills et al., 2017; Rivera and Tilcsik 2016; Tilcsik 2011), and theories of employer evaluations of interrupted and atypical careers (e.g., Leung, 2014; Pedulla, 2016; 2020), we propose that employers will be less likely to interview ex-founders due to negative stereotypes they attach to founders returning to wage work. However, because these penalties derive not from aspects of the founding history but solely from the negative stereotypes that employers attach to ex-founders, we also propose that the penalty will be diminished when such stereotypes are attenuated. In particular, building upon research about the contextual influence of gender stereotypes (e.g., Blair 2002; Blair, Ma, and Lenton 2001; Dasgupta and Greenwald 2001; Galperin et al., 2019; Heilman, 1984; Power, Murphy, and Coover 1996) and research

on the gendering of entrepreneurship (Kanze et al., 2018; Thébaud 2010), we hypothesize that female ex-founders will be less penalized than male ex-founders because positive stereotypes about gender and the perception of entrepreneurship as a male-typed activity will negate the influence of negative stereotypes.

The evidence, from a resume audit and a survey-based experiment of marketing professionals, supports our hypotheses. First, we find baseline evidence that employers avoid ex-founders, such that pursuing entrepreneurship translates into a 35 percent reduction in the likelihood of receiving an interview. The survey experiment confirmed that the penalty is partially motivated by employer's perception that ex-founders are less likely to fit and less willing or able to work at traditional firms. Critically, the evidence also shows that this baseline penalty is highly gendered, as female founders are nearly twice (13 percent vs. 7 percent) as likely to receive a callback as their male counterparts. This gender discrepancy resulted from employers sharing other evaluators biases, but to opposite effect. In plain language, because employers did not see women as equally entrepreneurial, they also did not perceive female ex-founders as exhibiting the same undesirable entrepreneurial traits they saw in male ex-founders. In summary, we find that, whereas women face systematic disadvantage in entrepreneurship, being less well represented than men and less likely to succeed (Thébaud 2015b; Thébaud and Sharkey 2016), this gender disparity reverses at the exit stage, presenting female founders with better recruitment chances.

This study makes a number of contributions, most notably by extending past research on entrepreneurial careers. Scholars have often argued that entrepreneurship offers a path for upward mobility (Birch 1997; Holzer 1996; Keister 2000; Kim et al. 2013; Mills 1951; Ruef et al. 2003), rewarding the decision to found a new venture with improved status or access to better advancement opportunities (Sørensen and Sharkey 2014). In addition, a growing line of research conceptualizes entrepreneurship as one stage in a career path, similar to other career choices (e.g., Burton et al., 2017; Campbell 2013; Sørensen and Sharkey, 2014; Kacperczyk and Marx, 2016). Our study complements these research lines by attributing greater attention to the costs associated with an entrepreneurial career. Because entrepreneurship is intertwined with mobility decisions, as individuals move between paid

employment and entrepreneurship (Kacperczyk and Marx 2016), we document that employers will likely pay close attention to an entrepreneurship spell, when assessing the overall competence and fit of a job candidate as a prospective employee. The findings we present in this research highlight the potential costs of having been a founder – which tend to accumulate to individuals who seek re-employment, following an entrepreneurial stint. Whereas this cost at the critical stage of re-entry into paid employment has so far been unaccounted for, we suggest that taking prospective employers' evaluations into consideration is necessary to fully assess the sum of what individuals benefit and lose from the pursuit of entrepreneurship.

Our findings also extend research on hiring, in general, and the impact of interrupted and boundaryless careers, in particular, on employment outcomes (Arthur, 1994; Briscoe and Hall, 2006; Leung, 2014; Pedulla, 2016). By developing a theory about how employers evaluate ex-founders, we contribute to the growing line of research about the demand side of the hiring process (Bills et al., 2017; Rivera, 2011), in general, and employers' evaluations of erratic and non-standard employment, in particular (Pedulla 2016, Leung 2014; Galperin et al 2019, Weisshaar 2018). The findings presented here complement and extend the scope of this prior work by first, documenting employers' tendency to penalize job candidates with interrupted careers and second, by showing how such evaluations can shift depending on the demographic characteristics of the job candidate. Specifically, our study extends prior theories by providing evidence that typical employer penalties associated with interrupted careers cease to apply when commitment and fit are signaled through demographic attributes, such as is the case for gender. More generally, our findings offer critical boundary conditions to employer discounting on the basis of interrupted careers.

We also contribute novel findings to theories of female entrepreneurship, by shedding light on gender disparities at the exit stage, which has been neglected by past research. Although scholars have documented that the benefits of entrepreneurial careers appear to be unevenly distributed across women and men, with the former being more likely than the latter to enter and succeed in entrepreneurship (e.g., Kanze et al., 2018; Lee and Huang, 2018), conclusions about the degree of gender disadvantage in

entrepreneurship are incomplete because past research paid little attention to a significant stage of the entrepreneurial process: exit to paid employment. Our findings complement these past studies by documenting that entrepreneurship has no negative influence on women's chances in paid employment, even though entrepreneurial experience imposes penalties for male ex-founders. Thus, our study implies that the same biases that put women at a disadvantage at the point of entry may, paradoxically, benefit them at the point of exit, making women more insulated from the negative consequences of entrepreneurship. In this respect, our findings provide a welcome alternative and offer important evidence that entrepreneurship offers women a path for upward mobility and access to better advancement opportunities without limiting their future possibilities. As such, our study has implications for both private and public efforts to increase the number of women entering entrepreneurship.

Finally, our research has important policy implications. Scholars and policy makers have recently intensified efforts to increase the representation of minorities, including women, in entrepreneurship (e.g., U.S. Department of Commerce, Economics and Statistics Administration, 2010). These interventions aim to push women into entrepreneurship are coupled with a growing belief that increasing the number of women (and other minorities) in entrepreneurship can help reduce social inequality. At the same time, those policies are less attentive to the fact that individuals who are encouraged to enter may, at some point, need to return into paid employment. But to the extent that a spell of entrepreneurship will affect a person's ability to re-enter the labor market and thus their future advancement chances, more generally, it is important to conduct empirical analyses of how founding affects employer perception.

It is important to note that there are multiple limitations to these findings as our setting constrains the ability to generalize too broadly from the available evidence. Most obviously, we study the reception of mid-career individuals and there is reason to believe that younger and older applicants may be treated differently, and that entrepreneurship may function as a different signal for those individuals. Again, as we theorize a causal pathway where employer perceptions mediate the meaning of entrepreneurship, it is plausible that the age of the applicant similarly functions to influence understandings of the founder, and through this, the degree of penalty. Second, we limit our study to people in two occupations (HR and

marketing) because these are among the most common career paths for recent college graduates, but it is likely that our results are specific to similar occupations and would not translate to areas where employers may have more direct means of testing skills or in which the organizational culture is more similar to what one finds at start-ups. Third, we compare the outcomes for two qualified candidates (as evidenced by the high overall response rate), but our conclusions may strictly depend on the strength of a candidates' record or the tightness of labor markets. Fourth, many people acquire jobs through networking and personal contacts, but our findings are limited to how founding affects job seekers that lack those potentially moderating ties. Finally, we leverage the experimental survey as a way to further demonstrate our theory and test the plausibility of the mechanisms we hypothesize. However, given that the sample we use is not identical to that used in the audit study, our results should be interpreted with caution. Indeed, future studies may build on our research by designing interventions to allow for a more direct test of the mechanism at play.

Despite these limitations, the study offers contributions to a number of literatures as well as to contemporary dialogues around encouraging entrepreneurship. At the broadest level, our findings suggest that public efforts to encourage female entrepreneurship do not unintentionally handicap the very people they intend to promote. Instead, we find that women are more able than men to pursue entrepreneurship without harming their future employability. For scholars of careers and labor markets our findings provide evidence that employers discount job candidates with interrupted, entrepreneurial careers, but that such effect are far from uniform. Instead, they vary profoundly by gender. For scholars of entrepreneurship this provides necessary and missing evidence of the costs of founding, and helps clarify how biases against female founders persist, but also how their influence evolves. For scholars of inequality and stratification the findings reinforce prior claims that entrepreneurship can offer a vehicle for mobility and also provide evidence of how biases can invert to become benefits. At the same time, our results offer a cautionary tale to male founders and suggest that there is a large, and previously unrecognized, potential cost to founding that ought to be part of their calculus.

Table 1: Distribution of Applications by City (Study 1)

	Male Employee	Male Ex-Founder	Female Employee	Female Ex-Founder	All Applications	Callback Rate
<i>Highest Entrepreneurship</i>						
San Francisco	72 23.53%	52 16.94%	61 20.13%	60 19.54%	245 20.03%	10.0%
Boston	46 15.03%	50 16.29%	53 17.49%	52 16.94%	201 16.43%	14.4%
<i>High Entrepreneurship</i>						
Austin	16 5.23%	20 6.51%	12 3.96%	11 3.58%	59 4.82%	11.9%
San Diego	24 7.84%	15 4.89%	21 6.93%	21 6.84%	81 6.62%	21.0%
Washington, DC	29 9.48%	23 7.49%	26 8.58%	20 6.51%	98 8.01%	15.0%
<i>Moderate Entrepreneurship</i>						
Atlanta	17 5.56%	19 6.19%	27 8.91%	26 8.47%	89 7.28%	10.1%
Minneapolis	12 3.92%	21 6.84%	11 3.63%	16 5.21%	60 4.91%	11.7%
Philadelphia	27 8.82%	22 7.17%	16 5.28%	19 6.19%	84 6.87%	22.6%
<i>Low Entrepreneurship</i>						
Cleveland	16 5.23%	19 6.19%	12 3.96%	12 3.91%	59 4.82%	20.0%
Dallas	15 4.9%	19 6.19%	20 6.6%	28 9.12%	82 6.7%	9.8%
Detroit	21 6.86%	32 10.42%	25 8.25%	24 7.82%	102 8.34%	5.0%
Phoenix	11 3.59%	15 4.89%	19 6.27%	18 5.86%	63 5.15%	19.0%
Marketing Position	160	149	141	149	599	15.0%
HR Position	146	158	162	158	624	11.9%
Callback Rate	16%	7.17%	17.50%	13.70%	13.5%	
Total	306	307	303	307	1,223	

Table 2: Summary of Job Posting Variables

Job Posting Variables	Description
Experience Requested (years)	Years of work experience requested/preferred
Mention “Entrepreneurship”	Posting referenced a preference for “entrepreneurial” skills (1/0)
EEOC Requirement	Applicant required to complete EEOC form (1/0)
Request "Independent" Employees	Posting referenced a preference for “independence” or “self-motivated” applicants (1/0)
Request "Leaders"	Posting referenced a preference for evidence of “leadership” applicants (1/0)
Mention Generalism	Posting referenced a preference for “generalist” or broadly-skilled applicants (1/0)
Firm Age	Years since founding
Firm Size (ln)	# of employees (logged)
Application Order	1 st or 2 nd resume submitted for given posting
Occupation	HR or Marketing position
Entrepreneurial City	Job is in a city where entrepreneurship is common (1/0)
Application Timing	# of days after the job was first posted
Mention Salary	Did posting specify a salary or salary range? (1/0)
Industrial Sector	The firm’s principal business was in: Utilities, Healthcare, IT, etc...

Table 3: Correlation Matrix (Study 1)

Variables	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)
(1) Interview Received	1												
(2) Ex-Founder	-0.0623	1											
(3) Female Candidate	0.0773	0.1492	1										
(4) Experience Requested (years)	-0.0557	-0.0019	-0.0001	1									
(5) Job Position (Midlevel, Senior)	-0.0399	-0.0028	-0.0242	0.3419	1								
(6) EEOC Requirement	-0.0336	-0.0027	0.0084	0.0416	-0.017	1							
(7) Request "Independent" Employees	0.0159	-0.009	0.003	-0.0135	-0.028	0.0392	1						
(8) Request "Leaders"	0.0121	-0.0095	0.0008	0.2144	0.0491	0.0259	0.0858	1					
(9) Mention "Entrepreneurship"	-0.0166	-0.0083	0.0013	0.0769	0.0443	0.0524	0.1043	0.1815	1				
(10) Mention Generalism	0.0068	-0.0032	0.0011	0.015	0.0452	-0.0134	0.042	-0.0194	-0.0337	1			
(11) Mention Salary	0.0108	0.0055	0.0243	0.0321	0.0096	0.0101	-0.0249	-0.0449	-0.0338	0.1212	1		
(12) Firm Age	-0.0195	-0.0249	-0.0153	0.0095	0.0087	-0.0231	0.0028	0.009	-0.0365	0.021	0.037	1	
(13) Firm Size (ln)	-0.0553	-0.0023	-0.0906	0.0476	0.0052	0.0208	-0.0243	0.0497	-0.0153	0.0284	-0.026	0.2449	1
(14) Sector	-0.0259	0.0066	0.0411	-0.023	-0.028	-0.0824	-0.0078	-0.0376	-0.0669	0.0267	0.0563	0.0521	-0.0519
(15) Application Order	-0.007	-0.0408	0.0237	0.0036	0.0014	-0.0026	0.0001	0.0032	0.0021	0	-0.0043	0.0166	0.014
(16) Occupation (HR or Marketing)	-0.0302	0.0041	0.0338	0.0644	0.0453	-0.0469	-0.0246	0.0035	-0.0768	0.1345	0.3232	0.1134	0.0228
(17) Entrepreneurial City	-0.0214	-0.0206	-0.0089	0.0629	-0.0335	0.0815	0.1344	0.219	0.1227	-0.0586	0.0143	-0.0469	0.035
(18) Fiscal Quarter	0.0223	0.0294	0.0096	0.0148	0.0051	-0.0057	-0.0916	-0.0221	-0.0067	0.0337	0.1126	-0.0124	-0.1111
(19) Application Timing (days from posting)	-0.0121	-0.0299	0.0059	0.005	-0.0186	0.0285	-0.0167	-0.0154	-0.0318	-0.0106	-0.0131	0.0447	0.0259
Variables	(14)	(15)	(16)	(17)	(18)								
(1) Interview Received													
(2) Ex-Founder													
(3) Female Candidate													
(4) Experience Requested (years)													
(5) Job Position (Midlevel, Senior)													
(6) EEOC Requirement													
(7) Request "Independent" Employees													
(8) Request "Leaders"													
(9) Mention "Entrepreneurship"													
(10) Mention Generalism													
(11) Mention Salary													
(12) Firm Age													
(13) Firm Size (ln)													
(14) Sector	1												
(15) Application Order	-0.0037	1											
(16) Occupation (HR or Marketing)	0.0577	-0.0097	1										
(17) Entrepreneurial City	-0.1013	0.0121	0.0585	1									
(18) Fiscal Quarter	0.0673	0.0015	0.1873	-0.0688	1								
(19) Application Timing (days from posting)	0.0012	0.2049	-0.0032	-0.0765	-0.0197								

Table 3: Correlation Matrix (Study 1) (continued)

Variable	Obs	Mean	Std.Dev.	Min	Max
Interview Received	1223	.136	.343	0	1
Ex-Founder	1223	.5	.5	0	1
Female Candidate	1223	.499	.5	0	1
Experience Requested (years)	1223	3.605	1.717	1	12.5
Job Position (Midlevel, Senior)	1223	1.0378	.190	1	2
EEOC Requirement	1223	.095	.293	0	1
Request "Independent" Employees	1223	.131	.337	0	1
Request "Leaders"	1223	.147	.354	0	1
Mention "Entrepreneurship"	1223	.05	.218	0	1
Mention Generalism	1223	.02	.144	0	1
Mention Salary	1223	.30	.459	0	1
Firm Age	1223	26.08	22.816	1	182
Firm Size (ln)	1223	5.311	2.1232	1	12.50
Sector	1223	3.813	.618	1	5
Application Order	1223	1.477	.5	1	2
Occupation (HR or Marketing)	1223	.51	.5	0	1
Entrepreneurial City	1223	1.757	1.117	0	3
Fiscal Quarter	1223	2.335	1.208	0	4
Application Timing (days from posting)	1223	4.469	3.836	1	31

Table 4: Logit Models of Entrepreneurial Experience and Interview (Study 1)

	(1)	(2)	(3)	(4)	(5)
Ex-Founder	-0.536*** (0.193)	-0.897*** (0.119)	-0.901*** (0.143)	-0.926*** (0.163)	-0.487*** (0.083)
Female Candidate		0.102 (0.126)	0.166 (0.104)	0.197** (0.095)	0.117** (0.052)
Ex-Founder * Female Cand.		0.621*** (0.198)	0.592*** (0.213)	0.607*** (0.183)	0.303*** (0.092)
Experience Requested (years)				-0.154*** (0.050)	-0.077*** (0.024)
Mention "Entrepreneurship"				-0.036 (0.278)	-0.046 (0.154)
EEOC Requirement				-0.680*** (0.254)	-0.353*** (0.123)
Request "Independent" Employees				0.163 (0.330)	0.114 (0.174)
Request "Leaders"				0.521** (0.252)	0.279** (0.139)
Application Order				-0.256*** (0.094)	-0.133*** (0.042)
Occupation				-0.366 (0.258)	-0.207 (0.138)
Entrepreneurial City				-0.879*** (0.079)	-0.481*** (0.038)
Application Timing				-0.028** (0.014)	-0.015* (0.008)
Mention Salary				0.148 (0.156)	0.096 (0.082)
Mention Generalism				0.884 (0.606)	0.480 (0.355)
Firm Age				-0.001 (0.003)	-0.001 (0.001)
Firm Size (ln)				-0.028 (0.048)	-0.018 (0.028)
Constant	-1.609*** (0.048)	-1.661*** (0.020)	-1.719** (0.826)	0.413 (1.242)	0.104 (0.601)
City Fixed Effect	No	No	Yes	Yes	Yes
Quarter Fixed Effect	No	No	Yes	Yes	Yes
Position Level Fixed Effect	No	No	Yes	Yes	Yes
Market Sector Fixed Effect	No	No	Yes	Yes	Yes
Observations	1,223	1,223	1,223	1,223	1,223
Model	Logit	Logit	Logit	Logit	Probit
Log Likelihood	-480.7	-477.0	-461.4	-446.3	-446.5

*** p<0.01, ** p<0.05, * p<0.1

Table 5: OLS Models of Recommendation to Hire (Study 2)

VARIABLES	Recommend	Recommend
	Full Sample	Full Sample
	(1)	(2)
Ex-Founder	-0.328** (0.159)	-0.765*** (0.220)
Female Candidate		0.230 (0.218)
Ex-Founder * Female Candidate		0.803** (0.312)
Constant	5.736*** (0.111)	5.628*** (0.149)
Observations	413	413
Adjusted R-Squared	0.010	0.062

*** p<0.01, ** p<0.05, * p<0.1

Table 6: OLS Models of Recommendation to Hire: Fit for Traditional Firm (Study 2)

VARIABLES PANEL A	Recommend Full Sample	Fit	Recommend
	(1)	(2)	(3)
Ex-Founder	-0.765*** (0.220)	-0.658*** (0.167)	-0.481** (0.212)
Female Candidate	0.230 (0.218)	-0.172 (0.165)	0.304 (0.206)
Founder * Female Candidate	0.803** (0.312)	0.528** (0.236)	0.575* (0.297)
Fit			0.432*** (0.062)
Constant	5.628*** (0.149)	4.858*** (0.113)	3.529*** (0.331)
Observations	413	413	413
R-squared	0.062	0.039	0.162

VARIABLES PANEL B	Recommend Female	Fit Female	Recommend Female
	(1)	(2)	(3)
Ex-Founder	0.038 (0.190)	-0.130 (0.160)	0.077 (0.185)
Fit			0.306*** (0.081)
Constant	5.859*** (0.137)	4.687*** (0.115)	4.426*** (0.400)
Observations	205	205	205
R-squared	0.000	0.003	0.067

VARIABLES PANEL C	Recommend Male	Fit Male	Recommend Male
	(1)	(2)	(3)
Ex-Founder	-0.765*** (0.246)	-0.658*** (0.173)	-0.411* (0.237)
Fit			0.538*** (0.092)
Constant	5.628*** (0.167)	4.858*** (0.117)	3.013*** (0.473)
Observations	208	208	208
R-squared	0.045	0.066	0.181

*** p<0.01, ** p<0.05, * p<0.1

Table 7: OLS Models of Recommendation to Hire: Dimensions of Fit (Study 2)

VARIABLES	Likely to Quit	Recommend	Easy to Manage	Recommend	Team Player	Recommend
	(1)	(2)	(3)	(4)	(5)	(6)
Ex-Founder	1.003*** (0.227)	-0.550** (0.220)	-0.642*** (0.152)	-0.443** (0.211)	-1.048*** (0.153)	0.242 (0.218)
Female Candidate	0.051 (0.225)	0.241 (0.212)	-0.220 (0.151)	0.341* (0.205)	0.253* (0.151)	0.104 (0.205)
Founder * Female Candidate	-0.610* (0.322)	0.672** (0.306)	0.377* (0.216)	0.614** (0.294)	0.805*** (0.217)	0.401 (0.298)
Likely to Quit		-0.214*** (0.047)				
Easy to Manage				0.502*** (0.067)		
Team Player						0.499*** (0.067)
Constant	3.292*** (0.153)	6.333*** (0.212)	4.947*** (0.103)	3.144*** (0.360)	4.858*** (0.103)	3.203*** (0.353)
Observations	413	413	413	413	413	413
R-squared	0.056	0.108	0.049	0.175	0.165	0.175

*** p<0.01, ** p<0.05, * p<0.1

Table 8: OLS Models of Recommendation to Hire: Fit for Recent Job (Study 2)

VARIABLES	Fit for Recent. Job	Fit for Recent Job Female	Fit for Recent Job Male
	(1)	(2)	(3)
Ex-Founder	0.114 (0.138)	-0.353** (0.145)	0.114 (0.131)
Female Candidate	0.008 (0.137)		
Founder * Female	-0.468** (0.196)		
Constant	4.770*** (0.093)	4.778*** (0.104)	4.770*** (0.089)
Observations	413	205	208
R-squared	0.030	0.028	0.004

Figure 1: Response to Entrepreneurship by Gender of Founder

Note: This graph includes data from 1,223 observations.

Appendix A. Table 1: Resume Randomization by Condition

Resume Element		<i>Male Founder</i>	<i>Male Non-Founder</i>	<i>Female Founder</i>	<i>Female Non-Founder</i>
1	Format	Randomized	Randomized	Randomized	Randomized
2	Address				
3	Last Name				
4	College				
5	Major				
6	1st Job				
7	2nd Job				
8	Skills				
9	3rd Job	Founder	HR Manager or Marketing Coordinator	Founder	HR Manager or Marketing Coordinator
10	First Name	Joe or Charlie	Joe or Charlie	Katie or Heidi	Katie or Heidi

Appendix A. Table 2: Logit Regressions of Job Interview (Study 1)

VARIABLES	(1)	(2)	(3)	(4)
	EEOC	Mechanisms	Manager Positions	Quality
Ex-Founder	-0.577*** (0.216)	0.165 (0.438)	-1.429*** (0.305)	-1.127*** (0.394)
Female Candidate	0.432** (0.174)	0.447*** (0.136)	-0.249 (0.307)	-0.480*** (0.139)
Ex-Founder * Female Candidate			1.471*** (0.465)	0.803*** (0.061)
EEOC Requirements * Female Candidate	0.155 (1.183)			
Experience Requested (years)	-0.153*** (0.051)	-0.153*** (0.056)	-0.227** (0.109)	-0.271*** (0.103)
Mention “Entrepreneurship”	-0.016 (0.275)	-0.163 (0.311)	0.481 (0.512)	-0.026 (0.272)
EEOC Requirement	-0.780 (0.795)	-0.666*** (0.255)	-0.730* (0.441)	-0.683** (0.276)
Request “Independent” Employees	0.162 (0.323)	0.389 (0.268)	0.657 (0.662)	0.160 (0.326)
Request “Leaders”	0.506** (0.240)	0.382 (0.263)	0.041 (0.301)	0.519* (0.273)
Application Order	-0.246** (0.097)	-0.268** (0.105)	-0.559*** (0.136)	-0.253** (0.100)
Occupation (HR or Marketing)	-0.370 (0.261)	-0.376 (0.254)	0.028 (0.441)	-0.334 (0.233)
Entrepreneurial City	-0.886*** (0.072)	-0.924*** (0.086)	-0.591** (0.258)	-0.858*** (0.064)
Application Timing (days from posting)	-0.029** (0.014)	-0.026** (0.013)	0.055 (0.056)	-0.027** (0.013)
Firm Size (ln)	-0.026 (0.048)	0.011 (0.057)	-0.148* (0.087)	-0.029 (0.050)
Mentions Salary	0.146 (0.161)	0.156 (0.159)	0.122 (0.338)	0.163 (0.171)
Mentions Generalism	0.874 (0.581)	0.844 (0.556)	0.963 (1.226)	0.963 (0.630)
Firm Age	-0.001 (0.003)	0.002 (0.002)	0.012 (0.009)	-0.001 (0.002)
Ex-Founder * Mention “Entrepreneurship”		0.192*** (0.052)		

Appendix A. Table 2: Logit Regressions of Job Interview (Study 1) (continued)

VARIABLES	(1)	(2)	(3)	(4)
	EEOC	Mechanisms	Manager Positions	Quality
Ex-Founder * Request "Leaders"		0.337* (0.182)		
Ex-Founder * Request "Independent" Employees		-0.642 (0.431)		
Ex-Founder * Firm Age		-0.011** (0.005)		
Ex-Founder * Firm Size (ln)		-0.089 (0.064)		
Ex-Founder * Female Candidate * Exp. Requested				-0.014 (0.017)
Ex-Founder * Exp. Requested				0.057 (0.142)
Female Candidate * Exp. Requested				0.211*** (0.073)
Constant	0.336 (1.384)	0.078 (1.434)	2.390 (1.771)	-0.446 (0.540)
City Fixed Effect	Yes	Yes	Yes	Yes
Fiscal Quarter Fixed Effect	Yes	Yes	Yes	Yes
Position Level Fixed Effect	Yes	Yes	Yes	Yes
Market Sector Fixed Effect	Yes	Yes	Yes	Yes
Observations	1,223	1,223	366	1,223
Log Likelihood	-447.7	-445.1	-135	-446.3

*** p<0.01, ** p<0.05, * p<0.1

Appendix A. Table 3: OLS Models of Recommendation to Hire: Alternative Explanations (Study 2)

VARIABLES	Perceived Skill	Recommend Skill	Perceived Failure	Recommend Perceived Failure	Remember Founder Full Sample
	(1)	(2)	(3)	(4)	(5)
Ex-Founder	0.028 (0.139)	-0.780*** (0.208)	0.460** (0.220)	-0.735*** (0.221)	
Female Candidate	-0.053 (0.138)	0.258 (0.205)	0.119 (0.218)	0.238 (0.217)	0.491* (0.285)
Founder * Female	-0.193 (0.197)	0.905*** (0.295)	-0.240 (0.312)	0.787** (0.312)	
Perceived Skill		0.528*** (0.074)			
Perceived Failure				-0.066 (0.049)	
Constant	4.982*** (0.094)	2.997*** (0.393)	4.770*** (0.093)	4.778*** (0.104)	-0.148 (0.206)
Observations	413	413	413	413	413
Model	OLS	OLS	OLS	OLS	Logit
R-squared	0.009	0.167	0.030	0.028	
Log Likelihood					-137.5

*** p<0.01, ** p<0.05, * p<0.1

Appendix A. Figure 1: Examples of a Typical HR Submission

Male Control (Employee)

Joe Robinson	
829 NE Jackson St., Minneapolis MN 55413 jrobinson022@gmail.com (612)-568-4963	
SUMMARY	
Dynamic Human Resource professional with broad experience in all areas of HR	
<ul style="list-style-type: none">Extensive experience in performance management, change management, and HR consultancy.Professional expertise with excellent interpersonal and communication skills, an outstanding project management competence	
EXPERIENCE	
Lola Red PR (Minneapolis, MN)	May 2014 to present
HR Manager	
<ul style="list-style-type: none">Managed a team of 8 responsible for delivering professional service to managers and employees across a challenging and demanding client group. Provided advice on multiple HR issues and ensured the timely delivery of all HR related projects for designated client groups.Managed appraisal and performance issues, ensuring consistent application of the appraisal framework and the identification of training needs.Provided assistance with employee training system that addresses clients' training needs including training needs assessment, management development, the measurement of training impact, and training transfer.Effectively managed all disciplinary and grievance issues relating to absence, performance, in consultation with line managers and ensuring employment law is followed. Contributed to the continuing development of the HR policies and procedures by reviewing all significant changes to the policies and procedures.	
Best Buy (Minneapolis, MN)	June 2012 to May 2014
HR Coordinator	
<ul style="list-style-type: none">Developed and administrated corporate and human resource policies and procedures, including administration of all benefit plans and programs.Coached employees and counsel management in conflict resolution, motivation, goal setting, performance appraisals and staff development.Increased efficiency by creating and implementing processes, procedures, and eventual best practices to effectively support to the sales, project management, programming, and design departments across a number of offices creating performance metrics and accountability.Charged with recruitment, performance management, staff development, employee relations and policy implementations.	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
HR Analyst	
<ul style="list-style-type: none">Revised job descriptions across all levels and categories. "Shadowed" and interviewed employees to construct an accurate picture of the duties and skills required for each position.Developed comprehensive training programs and seminars, which were delivered to supervisors, technical employees, and management personnel.	

Male Treatment (Founder)

Charlie Lewis	
5002 33rd Ave. S. Minneapolis, MN 55417 founder.coo@fusionsolved.com (612)-930-0436	
Summary	
HR services professional with excellent record for designing, developing, initiating, and generating a number of new HR solutions. Served as project manager and business founder for consulting teams, providing HR services to client groups across different industries and in all HR functions.	
Experience	
Fusion Digital Solutions Inc.	Minneapolis, MN
<i>Co-Founder and Chief Operating Officer</i> 2014 - present	
<ul style="list-style-type: none">Successfully established and built a 10-person company providing HR solutions to high-profile clients. Built solid reputation for delivering high-quality, cost-effective solutions for customized products to address client needs. Offered proactive support to a diverse set of clients over the full range of HR tasks.Managed all aspects of business and project development from initial scoping through strategic and fiscal planning to final testing and delivery.Provided direction and guidance during organizational changes, balancing the expectations and needs of the client organizations, their employees, stakeholders, and customers.Established standardized personnel selection procedures, including interviewing, testing and reference, and background checking in a variety of client companies across different industries.	
Target	Minneapolis, MN
<i>Project Manager – HR Division</i> 2012 - 2014	
<ul style="list-style-type: none">Devised a recruitment and talent management program spanning the full range of expert and management positions.Recognized throughout the organization for creative development of effective recruitment practices, including innovative interviewing techniques to identify and hire appropriate candidates, which have served to foster employee retention.Led organizational assessment, diagnosis and the implementation of an organizational change initiative. Identified initiatives, made recommendations and trained managers and directors on succession planning and competency-based selection tools.Assisted in developing new HR system, which tracked all employee information, performance, pay increases, vacation, sick time, and functional responsibilities.	
Target	Minneapolis, MN
<i>HR Generalist</i> 2010 - 2012	
<ul style="list-style-type: none">Identified initiatives, made recommendations and trained managers and directors on succession planning and competency-based selection tools.Provided HR services such as staffing, HR development, and leadership coachingLed several initiatives and projects across different HR functions, including training, leadership, planning, and performance management.Effectively managed employee communication. Responsible for change and integration activities including employee engagement events, implementation of psychometric assessments, redundancy and redeployment programs, harmonization of benefits, terms and conditions.	

Note: The above depicts examples of the two male conditions, and the two different application styles and names. It is important to reiterate that the name/application combination was randomly assigned, so “Charlie Lewis” was equally likely to be a founder and “Joe Robinson” the employee for any given application.

Appendix A. Figure 2: Baseline Entrepreneurial Penalty

Note: This graph combines the results for male and female candidates, n=1,223.

Appendix A. Figure 3: The Effect of Gender of Entrepreneurship Penalty (Study 1)

Appendix A. Figure 4: Minneapolis Male Marketing Resumes (Study 1 & 2)

Male Control (Employee)

Joe Robinson	
829 NE Jackson St., Minneapolis MN 55413 jrobinson022@gmail.com (612)-568-4963	
SUMMARY	
Creative Marketing professional in brand marketing and internet marketing. <ul style="list-style-type: none">Developer of creative online branding for several key brands.Excellent communicator in brand and customer marketing portfolios.	
MARKETING EXPERIENCE	
Lola Red PR (Minneapolis, MN)	May 2014 to present
Director of Product Marketing <ul style="list-style-type: none">Managed a team of 8 to develop corporate branding standards for print and online media. Responsible for developing solutions for customized products addressing client needs.Coordinated all aspects of business and project development from initial scoping through strategic and fiscal planning to final testing and delivery.Steered full lifecycle product marketing management, developing and implementing marketing solutions.Administered marketing campaigns to promote client products. Developed innovative tools to help client analyze and define target markets for new product features and devise go-to-market strategies for national and international markets.	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Coordinator <ul style="list-style-type: none">Helped develop regional B2B marketing strategiesManaged a \$1m marketing campaign that resulted in 7% increase in regional salesDeveloped customized products to target client needs and increase salesCreated marketing collateral including catalogs, direct mail pieces and advertisementsManaged the company's online presence on LinkedIn, Facebook, and Twitter	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Assistant <ul style="list-style-type: none">Wrote a weekly report on company-wide marketing activities for Regional VPsServed as a liaison between company and outside vendorsDesigned and launched regular e-mail campaignsDeveloped strategies and coordinated activities for social media campaigns	
EDUCATION	
University of Illinois, Urbana-Champaign	June 2010
B.A. with Major in Marketing	
SKILLS	
Expertise in Tableau, Mailchimp, Hubspot, Marketo, Wordstream, Hootsuite, Google Analytics and various SEO tools. Familiarity with HTML, Java, CSS, and Shopify.	
INTERESTS	
Tennis, travel, cooking.	

Male Treatment (Founder)

Joe Robinson	
829 NE Jackson St., Minneapolis MN 55413 robinson@fusionsolved.com (612)-930-0436	
SUMMARY	
Creative Marketing professional in brand marketing and internet marketing. <ul style="list-style-type: none">Developer of creative online branding for several key brands.Excellent communicator in brand and customer marketing portfolios.	
MARKETING EXPERIENCE	
Fusion Digital Solutions Inc. (Minneapolis, MN)	May 2014 to present
Co-Founder and Chief Marketing Officer <ul style="list-style-type: none">Successfully established and built a 10-person company providing marketing solutions to high-profile clients. Built solid reputation for delivering high-quality, cost-effective solutions for customized products to address client needs. Steered full lifecycle product marketing management, developing and implementing marketing solutions.Managed all aspects of business and project development from initial scoping through strategic and fiscal planning to final testing and delivery.Administered marketing campaigns to promote client products. Developed innovative tools to help clients analyze and define target markets for new product features and devise go-to-market strategies for national and international markets.	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Coordinator <ul style="list-style-type: none">Helped develop regional B2B marketing strategiesManaged a \$1m marketing campaign that resulted in 7% increase in regional salesDeveloped customized products to target client needs and increase salesCreated marketing collateral including catalogs, direct mail pieces and advertisementsFacilitated company's online presence on LinkedIn and Twitter	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Assistant <ul style="list-style-type: none">Wrote a weekly report on company-wide marketing activities for Regional VPsServed as a liaison between company and outside vendorsDesigned and launched regular e-mail campaignsDeveloped strategies and coordinated activities for social media campaigns	
EDUCATION	
University of Illinois, Urbana-Champaign	June 2010
B.A. with Major in Marketing	
SKILLS	
Expertise in Tableau, Mailchimp, Hubspot, Marketo, Wordstream, Hootsuite, Google Analytics and various SEO tools. Familiarity with HTML, Java, CSS, and Shopify.	
INTERESTS	
Tennis, travel, cooking.	

Appendix A. Figure 5: Minneapolis Female Marketing Resumes (Study 1 & 2)

Female Control (Employee)

Katie Robinson	
829 NE Jackson St., Minneapolis MN 55413 rbnsn1990@gmail.com (612)-568-4963	
SUMMARY	
Creative Marketing professional in brand marketing and internet marketing. <ul style="list-style-type: none">Developer of creative online branding for several key brands.Excellent communicator in brand and customer marketing portfolios.	
MARKETING EXPERIENCE	
Lola Red PR (Minneapolis, MN)	May 2014 to present
Director of Product Marketing <ul style="list-style-type: none">Managed a team of 8 to develop corporate branding standards for print and online media. Responsible for developing solutions for customized products addressing client needs.Coordinated all aspects of business and project development from initial scoping through strategic and fiscal planning to final testing and delivery.Steered full lifecycle product marketing management, developing and implementing marketing solutions.Administered marketing campaigns to promote client products. Developed innovative tools to help client analyze and define target markets for new product features and devise go-to-market strategies for national and international markets.	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Coordinator <ul style="list-style-type: none">Helped develop regional B2B marketing strategiesManaged a \$1m marketing campaign that resulted in 7% increase in regional salesDeveloped customized products to target client needs and increase salesCreated marketing collateral including catalogs, direct mail pieces and advertisementsManaged the company's online presence on LinkedIn, Facebook, and Twitter	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Assistant <ul style="list-style-type: none">Wrote a weekly report on company-wide marketing activities for Regional VPsServed as a liaison between company and outside vendorsDesigned and launched regular e-mail campaignsDeveloped strategies and coordinated activities for social media campaigns	
EDUCATION	
University of Illinois, Urbana-Champaign	June 2010
B.A. with Major in Marketing	
SKILLS	
Expertise in Tableau, Mailchimp, Hubspot, Marketo, Wordstream, Hootsuite, Google Analytics and various SEO tools. Familiarity with HTML, Java, CSS, and Shopify.	
INTERESTS	
Baking, softball, soccer.	

Female Treatment (Founder)

Katie Robinson	
829 NE Jackson St., Minneapolis MN 55413 robinson@fusionsolved.com (612)-930-0436	
SUMMARY	
Creative Marketing professional in brand marketing and internet marketing. <ul style="list-style-type: none">Developer of creative online branding for several key brands.Excellent communicator in brand and customer marketing portfolios.	
MARKETING EXPERIENCE	
Fusion Digital Solutions Inc. (Minneapolis, MN)	May 2014 to present
Co-Founder and Chief Marketing Officer <ul style="list-style-type: none">Successfully established and built a 10-person company providing marketing solutions to high-profile clients. Built solid reputation for delivering high-quality, cost-effective solutions for customized products to address client needs. Steered full lifecycle product marketing management, developing and implementing marketing solutions.Managed all aspects of business and project development from initial scoping through strategic and fiscal planning to final testing and delivery.Administered marketing campaigns to promote client products. Developed innovative tools to help clients analyze and define target markets for new product features and devise go-to-market strategies for national and international markets.	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Coordinator <ul style="list-style-type: none">Helped develop regional B2B marketing strategiesManaged a \$1m marketing campaign that resulted in 7% increase in regional salesDeveloped customized products to target client needs and increase salesCreated marketing collateral including catalogs, direct mail pieces and advertisementsFacilitated company's online presence on LinkedIn and Twitter	
Best Buy (Minneapolis, MN)	June 2010 to May 2012
Marketing Assistant <ul style="list-style-type: none">Wrote a weekly report on company-wide marketing activities for Regional VPsServed as a liaison between company and outside vendorsDesigned and launched regular e-mail campaignsDeveloped strategies and coordinated activities for social media campaigns	
EDUCATION	
University of Illinois, Urbana-Champaign	June 2010
B.A. with Major in Marketing	
SKILLS	
Expertise in Tableau, Mailchimp, Hubspot, Marketo, Wordstream, Hootsuite, Google Analytics and various SEO tools. Familiarity with HTML, Java, CSS, and Shopify.	
INTERESTS	
Baking, softball, soccer.	

Appendix B, Figure 1: Survey Experiment – Instructions (Study 2)

INSTRUCTIONS

We are seeking help evaluating a marketing resume. You are about to see a real resume and then we will ask you a number of questions to help us evaluate the resume. Your input will help us greatly in understanding how people use resumes. Please answer to the best of your knowledge.

Appendix B: Table 1. Descriptive Statistics (Study 2)					
Variable	Obs	Mean	Std. Dev.	Min	Max
Recommend Hire	413	5.576	1.625	1	7
Fit for Traditional Firm	413	5.588	1.216	1	7
Likely to Quit	413	3.619	1.621	1	7
Easy to Manage	413	4.622	1.118	1	7
Team Player	413	5.753	1.177	1	7
Fit for Recent Job	413	5.709	1.003	1	7
Female Candidate	413	0.496	0.500	0	1
Ex-Founder	413	0.4866	0.500	0	1
Respondent Gender	413	1.530	0.499	0	1
Respondent Age	413	47.91	13.083	20	60
Job Experience	413	2.719	1.051	1	4
Education Level	413	2.225	0.587	1	3

Appendix B: Table 2. Correlation Matrix (Study 2)

Variable	1	2	3	4	5	6	7	8	9	10	11
Recommend Hire	1										
Misfit	0.3472	1									
Generalist	-0.254	-0.0593	1								
Failure	0.3564	0.4991	-0.0477	1							
De-Commitment	0.2471	0.3827	0.0157	0.4123	1						
Female Candidate	0.1845	0.0254	-0.061	-0.0285	-0.1132	1					
Ex-Founder	-0.1009	-0.1605	0.2047	-0.2042	-0.0657	0.0604	1				
Entrepreneurship Experience	-0.0604	-0.183	-0.2008	-0.1359	-0.1711	0.0029	-0.0736	1			
Respondent Age	0.1228	-0.0612	-0.0862	-0.0606	-0.0017	0.0137	-0.0715	0.1496	1		
Job Experience	0.0537	-0.0375	-0.1059	-0.0533	-0.0085	0.0303	-0.0671	0.0717	0.5814	1	
Education Level	0.077	0.1478	0.1811	0.1414	0.2067	-0.0591	0.0197	-0.3141	0.0064	-0.0278	1

Appendix B: Figure 1. The causal model underlying the relationship between the founding history and the willingness to hire

Appendix B: Figure 2. The causal model underlying the relationship between the founding history and the willingness to hire: Men

Appendix B: Figure 3. The causal model underlying the relationship between the founding history and the willingness to hire: Women

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